

# A Remote Sensing-Based Tool for Assessing Rainfall-Driven Hazards

Daniel B. Wright<sup>1\*</sup>

Ricardo Mantilla<sup>2</sup>

Christa D. Peters-Lidard<sup>3</sup>

<sup>1</sup>University of Wisconsin-Madison, Madison, Wisconsin

<sup>2</sup>University of Iowa, Iowa City, Iowa

<sup>3</sup>NASA Goddard Space Flight Center, Greenbelt, Maryland

[\\*danielb.wright@wisc.edu](mailto:danielb.wright@wisc.edu)

## Abstract

RainyDay is a Python-based platform that couples rainfall remote sensing data with Stochastic Storm Transposition (SST) for modeling rainfall-driven hazards such as floods and landslides. SST effectively lengthens the extreme rainfall record through temporal resampling and spatial transposition of observed storms from the surrounding region to create many extreme rainfall scenarios. Intensity-Duration-Frequency (IDF) curves are often used for hazard modeling but require long records to describe the distribution of rainfall depth and duration and do not provide information regarding rainfall space-time structure, limiting their usefulness to small scales. In contrast, RainyDay can be used for many hazard applications with 1-2 decades of data, and output rainfall scenarios incorporate detailed space-time structure from remote sensing. Thanks to global satellite coverage, RainyDay can be used in inaccessible areas and developing countries lacking ground measurements, though results are impacted by remote sensing errors. RainyDay can be useful for hazard modeling under nonstationary conditions.

Keywords: scenarios; extreme rainfall; remote sensing; floods; landslides; risk assessment

31   **Software Availability**

32   Name of Software: RainyDay Rainfall Hazard Modeling System

33   Developer: Daniel B. Wright

34   Contact: Daniel B. Wright; Address: Room 1269 Engineering Hall, 1415 Engineering Drive,  
35   Madison, WI 53706, USA; Email: [danielb.wright@wisc.edu](mailto:danielb.wright@wisc.edu)

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37   Year first available: 2015

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39   Required hardware and software: RainyDay requires Python 2.7 or newer (not tested with Python  
40   3.0 or higher) with Numpy and Scipy. The Netcdf4 and GDAL APIs and Python libraries are  
41   also required. RainyDay will run on Macintosh, Linux, and Windows machines with the proper  
42   APIs and Python libraries.

43

44   Cost: Free. RainyDay is currently available by request at  
45   <https://bitbucket.org/danielbwright/rainyday>. Open-source release under version 3.0 of the GNU  
46   General Public License (<http://www.gnu.org/licenses/gpl-3.0.en.html>) is planned, with  
47   unrestricted public access to the code repository.

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## 1. Introduction

Rainfall-driven hazards such as floods and landslides are the most common natural disasters worldwide, and amongst the most devastating. A growing number of computational hazard models are available to transform extreme rainfall inputs into hazard predictions, including distributed hydrologic models for the movement of water into and through river systems (e.g., Smith et al., 2004); hillslope stability and run-out models for landslide initiation and subsequent motion (e.g. Brenning, 2005 and Preisig and Zimmermann, 2010, respectively); and hydraulic models for flood wave propagation in channels and floodplains ( e.g., Horritt and Bates, 2002). These models have seen significant advances in recent decades, and have become key components in probabilistic hazard and risk assessment in fields such as natural catastrophe risk insurance, infrastructure design, and land-use planning. The hazard predictions produced by these models tend to be highly sensitive to the amount, timing, and spatial distribution of rainfall inputs. Unfortunately, progress on developing realistic rainfall inputs for probabilistic hazard and risk assessment has been relatively limited. This paper introduces RainyDay, a Python-based platform that addresses this shortcoming by coupling rainfall remote sensing data from satellites or other sources with a technique for temporal resampling and spatial transposition known as Stochastic Storm Transposition (SST) to generate highly realistic probabilistic rainfall scenarios.

Rainfall inputs for long-term hazard and risk assessment require a probabilistic description of three interrelated components: duration, intensity, and space-time structure. Efforts to jointly model these components are usually referred to as *rainfall frequency analysis*, a simple term that belies the complexity of the physical phenomena and analytical methods involved. The probability structure of the first two components, rainfall duration and intensity, has been a focus of research and application for decades (see U.S. Weather Bureau, 1958 and Yarnell, 1935 for early examples). These two components are strongly linked and together they determine the probability distribution of rainfall volume (or depth) at a point or over an area. The third component, space-time structure, describes the spatial and temporal variability of rainfall and is determined by storm size, horizontal velocity, and the temporal evolution of spatial rainfall coverage. Space-time structure can thus be understood as describing the “when” and “where” of extreme rainfall, whereas intensity and duration describe “how much.”

79  
80 Rainfall space-time structure can be an important hazard determinant. For example, a rainstorm  
81 that is short-lived and small in spatial extent may pose a significant flash flood threat in a narrow  
82 mountain valley or urban area, but may not represent a hazard on a larger river system.  
83 Conversely, a month-long rainy period could lead to flooding on a major river due to the gradual  
84 accumulation of water in soils, river channels, and reservoirs, but may never feature a short-lived  
85 “burst” of rainfall sufficiently intense to cause flash flooding at smaller scales. Similarly, a storm  
86 that covers a large area or passes over a series of valleys could lead to more widespread landslide  
87 or debris flow occurrences than a smaller or stationary storm. Rainfall space-time structure and  
88 its importance as a hazard trigger, therefore, must be understood within the context of the  
89 particular geography and scale in question. Due to its complexity, rainfall space-time structure  
90 has traditionally been less well understood than intensity and duration, and its representation in  
91 hazard modeling has been less sophisticated.

92  
93 The probability distribution of rainfall volume for a given duration is usually derived from rain  
94 gages and distilled into Intensity-Duration-Frequency (IDF) curves, such as those provided by  
95 the National Oceanic and Atmospheric Administration’s (NOAA) Atlas 14 (Bonnin et al., 2004).  
96 Long records (spanning many decades) are generally needed to define the extreme tail of such  
97 distributions. The challenge of measuring extreme rainfall over long time periods and over large  
98 areas using rain gages has hindered IDF estimation in many developed countries, while the lack  
99 of data in poor countries and in inaccessible terrain means that IDF estimation using such  
100 methods is virtually impossible in many locations. Furthermore, the ability to measure rainfall  
101 space-time structure at a high level of detail using dense networks of rain gages is nonexistent  
102 outside of a handful of wealthy cities and research-oriented observation networks.  
103 “Regionalization,”—the pooling of hazard information over a larger area in order to inform  
104 analyses at particular locations (see, e.g. Alexander, 1963 for an early discussion of rainfall  
105 regionalization and Stedinger et al., 1993 for a review)—has helped with IDF estimation using  
106 short records in areas where rain gage densities are moderate or high. These techniques offer  
107 little help, however, in parts of the world where rain gages are few or nonexistent, and do not  
108 offer a framework for incorporating rainfall space-time properties into hazard estimation. Even  
109 where long rainfall records do exist, nonstationarity due to climate change may mean that earlier



portions of the record are no longer representative of current or future IDF properties (e.g. Cheng and AghaKouchak, 2014).

Several techniques, which generally fall under the term of *design storm methods*, are used in long-term hazard estimation to link IDF properties to space-time structure for probabilistic flood hazard assessment (commonly referred to as *flood frequency analysis*). One central design storm concept is to link rainfall duration to rainfall intensity via a measure of flood response time, such as the time of concentration (e.g. McCuen, 1998). Another attempts to estimate area-averaged rainfall from point-scale rainfall estimates using area reduction factors (ARFs; U.S. Weather Bureau, 1958). Yet another uses dimensionless time distributions such as the family of U.S. Soil Conservation Service 24-hour rainfall distributions (e.g. McCuen, 1998). Each of these methods is highly empirical, laden with assumptions (see Wright et al., 2014a; Wright et al., 2014b; Wright et al., 2013), valid only in certain contexts, and often misunderstood or misused (K. Potter, personal communication, May 6, 2015).

SST explicitly links IDF properties rainfall space-time properties, providing certain advantages over design storm methods. Similar to other regionalization techniques, SST aims to effectively “lengthen” the period of record by using nearby observations, albeit using a fundamentally different approach involving temporal resampling and spatial transposition of rainstorms drawn from a catalog of observed rainfall events from the surrounding region. The inclusion of nearby storms at least partially addresses the difficulty of accurately estimating rainfall hazards using short records. SST can be used to estimate rainfall IDF properties and also to facilitate modeling of interactions of rainfall space-time structure with geographic features (such as hillslopes and river networks) at the appropriate spatial and temporal scales. It accomplishes this by generating large numbers of extreme rainfall “scenarios,” each of which has realistic rainfall structure based directly on observations.

Alexander (1963), Foufoula-Georgiou (1989), and Fontaine and Potter (1989) describe the general SST framework, while Wilson and Foufoula-Georgiou (1990) use the method for rainfall frequency analysis and Gupta (1972) and Franchini et al. (1996) use it for flood frequency analysis. In those days, however, the method was of limited practical use due to the lack of

detailed rainfall datasets with large areal coverage. Those studies also did not focus on the aspects of SST related to rainfall space-time structure nor its implications for hazard modeling.

The recent advent of satellite-based remote sensing provides a relatively low-cost means of measuring extreme rainfall over large parts of the globe at moderately high spatial and temporal resolution (30 minutes to 3 hours, 4 km to 25 km), while ground-based weather radar offers higher-resolution estimates (5-60 minutes, typically 1 to 4 km) over smaller regions. While the accuracy of rainfall remote sensing can be poor (particularly in cases of satellite-based estimates, e.g. Mehran and AghaKouchak, 2014; and mountainous regions, e.g. Nikolopoulos et al., 2013, Stampoulis et al., 2013), such data nonetheless offer unprecedented depictions of rainfall over large areas. This creates a variety of opportunities for hazards research and practice at various scales, ranging from forecasting and post-event analysis to long-term hazard assessment.

In the context of SST, the ongoing accumulation of remote sensing data to lengths of 10-20 years or more “unlocks” many of the as-yet unrealized opportunities offered by SST. Wright et al. (2013) demonstrated the coupling of SST with a 10-year high resolution radar rainfall dataset for IDF estimation, and the method was extended to flood frequency analysis for a small urban watershed using a distributed hydrologic model in Wright et al. (2014b). These two papers, along with Wright et al. (2014a) show that commonly-used design storm practices (ARFs, dimensionless time distributions) have serious shortcomings in representing the multi-scale space-time structure of extreme rainfall and critical interactions with of this structure with watershed and river network features. Wright et al. (2014b) also show that when SST is coupled with rainfall remote sensing data and a distributed hydrologic model, it can reproduce the role that this structure plays in determining multi-scale flood response. The RainyDay software described in this paper was developed to facilitate the use of SST in conjunction with rainfall remote sensing data.

Though SST was developed in the context of flood hazard estimation, it may prove useful for rainfall-triggered landslides and other mass movements, subject to the limited accuracy of remote sensing data in steep terrain and other limitations that will be discussed subsequently. Rainfall space-time structure governs the temporal distribution of rainfall volume onto individual

hillslopes, as well as the number of hillslopes subject to rainfall. In addition, steep landslide-prone terrain often has poorer rain gage coverage than lowland areas due to limited accessibility, suggesting that remote sensing rainfall estimates are potentially useful in such regions, particularly if improvements in accuracy can be realized (e.g. Shige et al., 2013).

Section 2 provides a description of the SST methodology. Section 3 discusses the specific implementation of SST in RainyDay, and some of the software’s important features. Section 4 provides sample results from RainyDay and sensitivity analyses using different input rainfall datasets for rainfall and flood frequency analysis in order to illustrate its capabilities and some of its limitations, including for flood frequency analysis in nonstationary conditions. Section 5 includes discussion and concluding remarks.

## **2. The SST Methodology**

In this section, we provide a step-by-step methodology for SST-based rainfall frequency analysis for a user-defined geographic “area of interest,”  $A$  of arbitrary shape. Higher-level description of software features is left to Section 3, but it merits mention that in RainyDay,  $A$  can be a single remote sensing pixel, a rectangular area containing multiple pixels, or a contiguous area defined by a user-supplied polygon shapefile.

The following five steps describe the SST methodology, as implemented in RainyDay:

1. Identify a geographic transposition domain  $A'$  that encompasses the area of interest  $A$ . One could confine  $A'$  to regions with homogeneous extreme rainfall properties, (e.g. flat areas far from large water bodies and topographic features). However, such homogeneity would likely be difficult to rigorously determine in practice and regardless, such strict interpretation is likely to be overly limiting. RainyDay offers several diagnostic aids, discussed in Section 3.3, that help the user to understand rainfall heterogeneity over the region  $A'$  and to improve the performance of the SST procedure in cases where rainfall heterogeneities do exist. Additional issues related to the selection of  $A'$  are explored in Section 4.3.

2. Identify the largest  $m$  temporally non-overlapping storms in  $A'$  from an  $n$ -year rainfall remote sensing dataset, in terms of rainfall accumulation of duration  $t$  and with the same size, shape, and orientation of  $A$ . For example, the principal axis of the Turkey River watershed in northeastern Iowa in the central United States is oriented roughly northwest-southeast and has an area of 4400 km<sup>2</sup>. In this case, the  $m$  storms are those associated with the  $m$  highest  $t$ -hour rainfall accumulations over an area of 4400 km<sup>2</sup> with the same size, shape, and orientation as the Turkey River watershed. We refer to this set of storms henceforth as a “storm catalog,” with the same geographic extent as  $A'$  and the same spatial and temporal resolution as the input rainfall data. We refer to the  $m$  storms in the storm catalog henceforth as “parent storms.” In RainyDay, the user can specify whether to exclude certain months (such as wintertime) from the storm catalog. Previous studies have shown that there can be low bias introduced in high-exceedance probability (i.e. frequent, low-intensity) events if  $m$  is small (e.g. Foufoula-Georgiou, 1989; Franchini et al., 1996; Wilson and Foufoula-Georgiou, 1990; see Wright et al., 2013 for a discussion). The sensitivity of SST results to the choice of  $m$  and  $A'$  is explored in detail in Section 4.3, but  $m \approx 10n$  generally minimizes the low bias for frequent events, and would likely be a good starting point for new analyses. Low exceedance probability (i.e. rare) events are less sensitive to the choice of  $m$  (see Section 4.3).

In RainyDay, duration  $t$  is a user-defined input, and as long as  $t$  is neither very short nor very long relative to the time scales of hazard response in  $A$ , subsequent hazard modeling results will be relatively insensitive to the chosen value. In this respect, the duration  $t$  in SST differs conceptually from design storm methods, in which hazard response is intrinsically sensitive to the user-specified duration, and this feature is indeed one of the chief advantages of SST over design storm methods for multi-scale flood hazard estimation (see Wright et al., 2014b for analysis and discussion). In the case of SST-based flood frequency analysis,  $t$  should be at least as long as the time of concentration and preferably somewhat longer.

3. Randomly generate an integer  $k$ , which represents a “number of storms per year.” In previous SST literature, the assumption was made that  $k$  follows a Poisson distribution

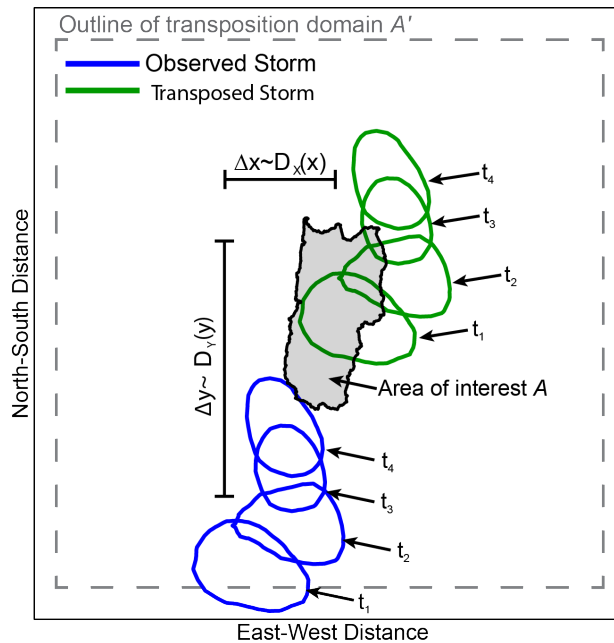
with a rate parameter  $\lambda$  storms per year. The  $m$  parent storms are selected such that an average of  $\lambda = m/n$  storms per year are included in the storm catalog. For example, if  $m = 100$  storms selected from a ten-year remote sensing dataset, then  $\lambda = 100/10 = 10.0$  storms per year. RainyDay will generate  $k$  using either the Poisson distribution or an empirical distribution, discussed in Section 3.3. If the Poisson distribution is selected, RainyDay will automatically calculate  $\lambda$  based on user-specified  $m$  and the length of the input dataset.

4. Randomly select  $k$  parent storms from the storm catalog. For each selected parent storm, transpose all rainfall fields associated with that storm by an east-west distance  $\Delta x$  and a north-south distance  $\Delta y$ , where  $\Delta x$  and  $\Delta y$  are drawn from the distributions  $D_X(x)$  and  $D_Y(y)$  which are bounded by the east-west and north-south extents of  $A'$ , respectively. The motion and structure of the parent storm is unaltered during transposition and only the location is changed. The distributions  $D_X(x)$  and  $D_Y(y)$  were taken to be uniform in Wright et al. (2013) and Wright et al. (2014b), but RainyDay offers additional options, described in Section 3.3. We illustrate this step schematically in Figure 1. For each of the  $k$  transposed storms, compute the resulting  $t$ -hour rainfall accumulation averaged over  $A$ .

Step 4 can be understood as temporal resampling and spatial transposition of observed storm events within a probabilistic framework to synthesize one year of heavy rainfall events over  $A'$  and, by extension, over  $A$ . RainyDay and previous SST efforts retain the largest (in terms of rainfall intensity) of the  $k$  events for subsequent steps and discard the  $k-1$  remaining events, though in principle these events could be retained. The single retained storm can be understood as a “synthetic” annual rainfall maximum, analogous to those annual rainfall maxima that are extracted from rain gage records for rainfall frequency analysis. It should be noted that these rainfall events do not form a continuous series, meaning that neither inter-storm periods nor the sequencing of the  $k$  storms are considered.

5. Repeat steps 3 and 4 a user-specified  $T_{max}$  number of times, in order to create  $T_{max}$  years of  $t$ -hour synthetic annual rainfall maxima for  $A$ . RainyDay then assigns each annual

maximum a rank  $i$  according to its rainfall intensity relative to all others. Each of these ranked maxima can then be assigned an annual exceedance probability  $p_e^i$  where  $p_e^i \equiv i/T_{max}$ . Exceedance probability  $p_e$  is the probability in a given year that an event of equal or greater intensity will occur. The “return period” or “recurrence interval”  $T_i$ , commonly used in hazard analysis, is simply  $T_i \equiv 1/p_e^i$ , so if  $T_{max} = 10^3$ , it is possible to directly infer exceedance probabilities of  $1.0 \geq p_e \geq 10^{-3}$  (recurrence intervals of  $1 \leq T_i \leq 10^3$ ). Each of these rainfall events can then serve as one datum of an empirical IDF estimate or as a rainfall scenario for hazard modeling.



**Figure 1: Depiction of SST procedure for a single storm consisting of four time intervals  $t_1...t_4$ . The blue ellipses illustrate the time evolution of an arbitrary rainfall isohyet derived from remote sensing observations, while the green ellipses show the time evolution of this same isohyet after transposition. Adapted from Wright et al. (2013).**

### 3. RainyDay Software

#### 3.1 Overview of Software

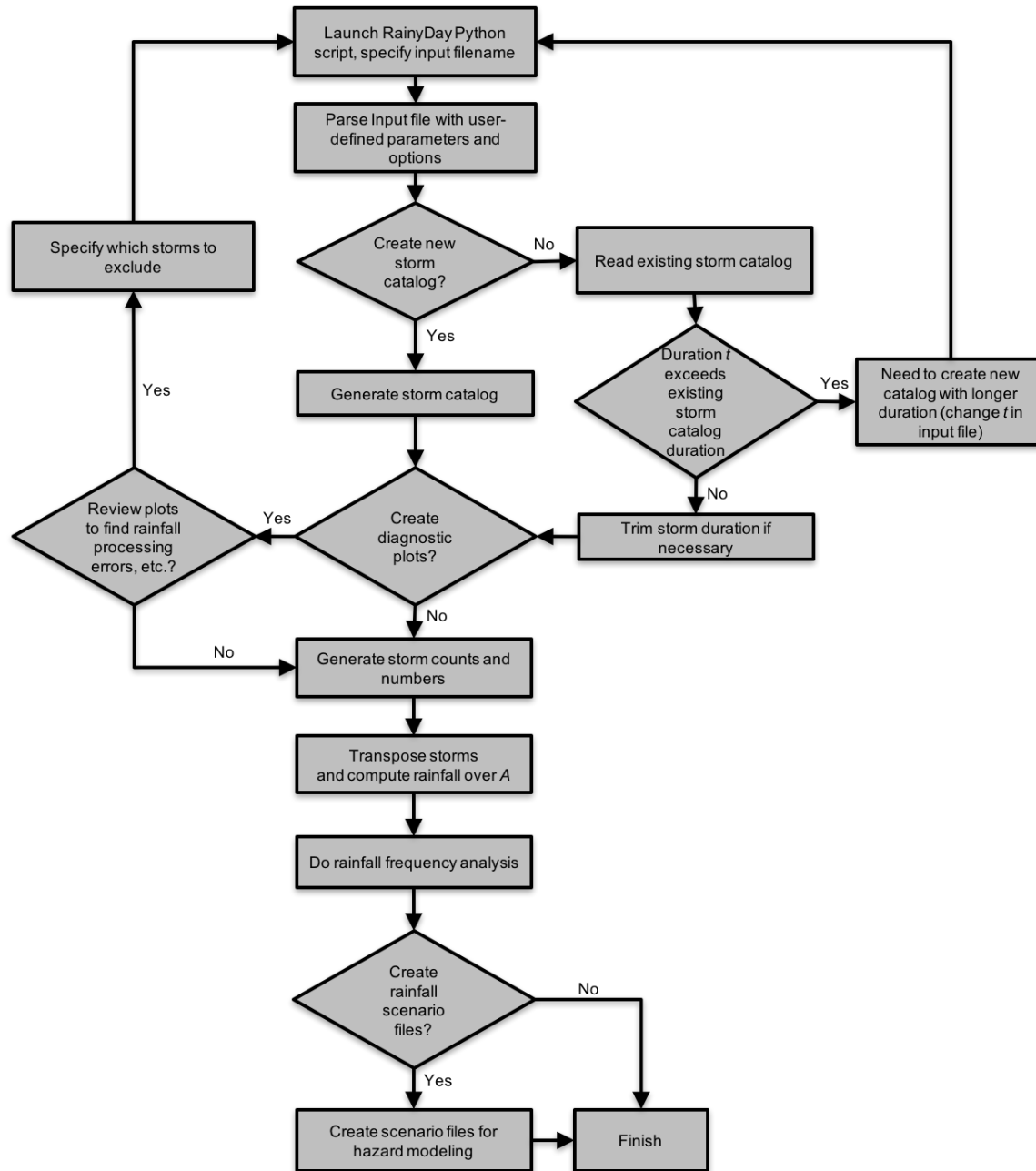
We wrote RainyDay to render SST more accessible and to streamline the code for speed and ease-of-use using Python. The majority of subroutines utilize the Scipy (Jones et al., 2011) and Numpy packages (Walt et al., 2011). To enhance speed, certain RainyDay subroutines call C

code through Scipy’s “weave” functionality  
(<http://docs.scipy.org/doc/scipy/reference/tutorial/weave.html>). Figure 2 shows a schematic of  
workflow in RainyDay.

While the ranking of rainfall events described in Step 5 of the SST methodology in Section 2 is  
based on rainfall intensity averaged over  $A$ , RainyDay will create NetCDF4 files  
(<http://www.unidata.ucar.edu/software/netcdf>) that contain the transposed rainfall scenarios with  
full depictions of rainfall space-time structure at the native spatial and temporal resolution of the  
input. This is an important feature because space-time structure, and not just average rainfall  
intensity over area  $A$  and duration  $t$ , is important in determining hazard response. For example,  
one rainfall scenario may produce a more severe flood response than another scenario, even if it  
has a lower overall average rainfall intensity over  $A$  and  $t$ , due to interactions with watershed  
features (see Section 3.2 of this paper for discussion and Wright et al., 2014b for analysis).

We will provide the RainyDay source code, examples, and user documentation upon request, and  
intend to release it under version 3 of the GNU General Public License  
(<http://www.gnu.org/copyleft/gpl.html>) once we have completed sufficient testing and  
documentation. The code is currently not parallelized, but shared-memory parallelization may  
be added in the future. Computational time is determined mainly by the size of the input dataset  
(record length  $n$ , input resolution, and geographic size of  $A$  and  $A'$ ), while other factors, such as  
 $m$ ,  $t$ ,  $T_{max}$ , and  $N$  can impact runtime. Computational speed, even without parallelization, is not  
prohibitive on a modern desktop or laptop computer (several seconds to several hours for typical  
configurations and input datasets).

To ensure accessibility for users inexperienced with Python, all of the necessary Python modules  
are supported within recent versions of the Anaconda Python distribution from Continuum  
Analytics (<https://store.continuum.io/cshop/anaconda>). The user must install NetCDF4 libraries  
and any requisite dependencies. If the user wishes to use shapefile functionality, necessary for  
defining  $A$  to be a shape other than a rectangle or a single rainfall pixel, the GDAL library  
(<http://www.gdal.org>) and any necessary dependencies must also be installed.



**Figure 2: Flow chart demonstrating the workflow of RainyDay.**

### 3.2 SST Internal Variability

In RainyDay, the user specifies  $N$ , the number of  $T_{max}$ -year long “ensemble members” to be generated. This enables the examination of “internal variability,” i.e. how much variation in rainfall intensity is possible for a given  $p_e$  for a given input rainfall dataset and set of user-defined parameters. For example, if the user specifies  $T_{max} = 10^3$  and  $N = 100$ , then there will be



100 intensity estimates for each  $p_e$  between 1.0 and  $10^{-3}$ . RainyDay will automatically generate a text file containing the results of this rainfall frequency analysis, including the rainfall mean, minimum, and maximum (or, optionally, a quantile interval) for each  $p_e$ , computed from the  $N$  ensemble members.

If the scenarios generated by RainyDay are fed through a hazard model, then the ensemble spread will propagate through to generate ensemble hazard estimates. A useful and interesting feature of SST and RainyDay that is not examined in this paper, but is discussed at length in Wright et al. (2014b), is that the exceedance probability of rainfall and of subsequent hazards can be decoupled using SST, particularly if some realistic scheme is used to account for the initial conditions in  $A$  (such as soil moisture or baseflow). Consider the example where  $N=1$  and  $10^3$  rainfall scenarios ( $T_{max}=10^3$ ) are created as input to a distributed flood hydrologic model. One of these rainfall scenarios has  $p_e=0.01$  (in terms of watershed-average  $t$ -hour rainfall depth over an area  $A$ ). Even if initial conditions are kept constant across all  $T_{max}$  simulations, the  $p_e$  of the peak discharge or volume predicted by the model for this particular scenario need not be equal to 0.01, since the space-time structure of the rainfall scenario and its interactions with watershed and river network features can dampen or magnify the flood severity. If variability in initial conditions within the hazard model are considered, this dampening or magnification effect can be even greater. This property of SST contrasts with design storm methods, which typically assume a 1:1 relationship between the  $p_e$  of rainfall and the resulting hazard, though variability in initial conditions could in principle be used with design storm approaches to produce some degree of “decoupling” of rainfall and hazard  $p_e$ . Setting  $N \geq 2$  allows for examination of differences in hazard  $p_e$  for a given rainfall  $p_e$ , or vice versa, which could lead the way to more detailed examination of the role of rainfall space-time structure (see Wright et al. 2014b) or initial conditions in probabilistic hazard estimation. RainyDay provides one simple scheme for creating variability in initial conditions, described in Section 3.5.

It should be pointed out that the ensemble spread generated in RainyDay is not completely comparable to the confidence intervals of more traditional rainfall or flood frequency analyses. The latter show statistical uncertainty associated with parameter estimation, which can be derived in different ways (e.g. bootstrapping, profile likelihood, etc.). Therefore, it might not be

reasonable to expect that the uncertainty ranges produced by RainyDay to be comparable to the confidence intervals of other IDF estimates. Like nearly all frequency analyses and IDF estimation methods, the ensemble spread generated by RainyDay does not consider measurement error, which, as mentioned previously, can be substantial. Since the ensemble spread is for a given set of user-defined values such as  $A'$  or  $m$ , it does not consider uncertainty associated with these choices. Analyses in Section 4.3 show how such uncertainties can be assessed, but fundamentally this requires manipulating the size or composition of the storm catalog through the choice of user-defined values, necessitating multiple distinct runs of RainyDay.

Ensemble spread is shown throughout Section 4 to illustrate various aspects of SST-based rainfall and flood frequency analysis. If the user is only interested in examining internal variability of SST-based rainfall IDF, then the number of ensemble members can be large (e.g.  $N \geq 100$ ). If the user wishes to perform hazard simulations, however,  $N$  should be selected with consideration of the computational cost associated with large numbers of simulations, which can be substantial depending on the particular hazard model. To help manage the number of simulations required, the user can specify a rainfall return period threshold, below which output scenarios will not be created. For example, if the user specifies a 5-year threshold, no rainfall scenarios with a rainfall depth less than the 5-year return period depth will be written, which reduces the number of hazard simulations by 80% for a given value of  $N$  while still retaining the most extreme scenarios.

### 3.3 *Rainfall Heterogeneity and Non-Uniform Spatial Transposition*

A common criticism of SST is that its validity is restricted to regions with homogenous extreme rainfall properties. As previously mentioned, depending on how rigidly this criterion is enforced, the method would be limited to small, flat regions far from topographic features, water bodies, etc. It is unclear how homogeneity would be determined, particularly with the paucity of extreme rainfall data in most regions. Instead, steps can be taken to use SST in more varied geophysical settings. Regardless of the setting, the selection of  $A'$  requires an understanding of regional rainfall patterns and of the intrinsic assumptions of SST. Though more work is needed to understand the geographic limits of the applicability of RainyDay in complex terrain, the work of England et al. (2014) provides an example of SST in complex terrain.

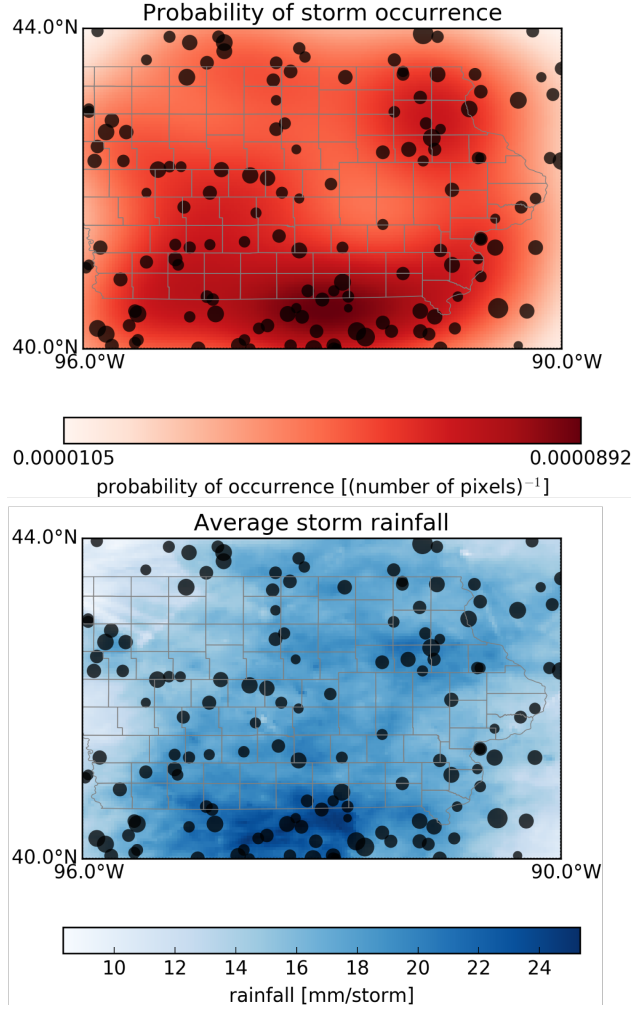
RainyDay provides several tools to help understand the issue of rainfall heterogeneity, and, to some extent, to mitigate it. First, RainyDay produces a map showing the location of the rainfall centroids for all storms in the storm catalog, overlaid on a smoothed field of the spatial probability of storm occurrence within  $A'$ . This spatial probability of occurrence map is generated by applying a two-dimensional Gaussian kernel smoother to the  $(x,y)$  locations of the rainfall centroids for all the storms in the storm catalog. This smoothed field is then normalized such that the sum of all grid cells is 1.0, thus creating a two-dimensional probability density function of storm occurrence. A second plot shows these rainfall centroids overlaid with the average rainfall per storm across  $A'$ . These diagnostic plots assist in understanding regional variations in storm occurrences and rainfall over  $A'$ . Examples of these diagnostic plots for a region  $A'$  encompassing most of the state of Iowa in the central United States are shown in Figure 3. The top panel suggests that storms are somewhat more frequent in the southernmost third or so of the transposition domain (top panel), along with slightly elevated activity in the northeast quadrant. The bottom panel shows somewhat higher average storm rainfall in these two areas. Caution should be taken when drawing firm conclusions from these diagnostic plots, however, since rainfall heterogeneities evident in both storm occurrences and average storm rainfall may be the result of spatial biases in rainfall remote sensing estimates or of randomness in the climate system over the relatively short remote sensing record, rather than from “true” heterogeneity in the underlying rainfall hydroclimate.

Additional optional diagnostic outputs include static and animated rainfall maps for each storm in the storm catalog (not shown). These storm rainfall maps are useful for diagnosing “bad data,” particularly in rainfall datasets that use ground-based weather radar contaminated by radar beam blockage and other unrealistic artifacts. RainyDay allows for the exclusion of user-identified storm periods from subsequent analysis, though anomalous periods must be identified by the user (i.e. no automatic data quality checking is provided).

The two-dimensional density function of spatial storm probability of storm occurrence can optionally be used as the basis for non-uniform spatial transposition (providing the  $D_X(x)$  and  $D_Y(y)$  described in Step 4 and Figure 1 in Section 2) so that the spatial distribution of storm

occurrences will be preserved between the input data and output rainfall scenarios and IDF estimates. Section 4.3 examines the impact of this optional feature on results for the Iowa study region, along with potential implications.

It is important to note that this approach only addresses the spatial heterogeneity of storm occurrences, not of spatial variations in the climatology of rainfall intensity (due to topography or other factors). For example, if  $A'$  contains both a flat plain and an adjacent mountain range, the probability of storm occurrence will vary across  $A'$ . This variation will be captured in the two-dimensional density function of spatial storm probability and, using the optional non-uniform spatial transposition scheme, will be reflected in RainyDay outputs. In this example, rainfall intensity from these storms will also vary according to the underlying topography. The current transposition scheme in RainyDay cannot explicitly account for this intensity variation. This is likely to be a serious constraint in many regions.



**Figure 3: Example of diagnostic plots produced by RainyDay for 24-hour duration rainfall from the Stage IV rainfall dataset (described in Section 4.1) over a region encompassing the state of Iowa, United States. Top: shading indicates spatial probability of storm occurrence. Bottom: shading indicates the average rainfall per storm from the same storm catalog. Black dots show the rainfall centroids for each storm in the storm catalog. Dot size in both panels indicates relative rainfall storm total rainfall depth. Key RainyDay parameters:  $m=150$  storms,  $A'=[40^\circ \text{ to } 44^\circ \text{ N}, 90^\circ \text{ to } 96^\circ \text{ W}]$ .  $A$  is a single Stage IV rainfall pixel (approximately  $16 \text{ km}^2$ ),  $T_{\max}=1000$ , and  $t=24$  hours.**

### 3.4 Empirical Temporal Resampling

As mentioned in Step 3 of the SST procedure described in Section 2, previous SST work has employed the assumption that the annual number of storm counts follows a Poisson distribution, which in turn serves as the basis for the temporal resampling of storms (i.e. for generating the number of storms per year  $k$  that will be spatially transposed). RainyDay supports Poisson-based resampling, but also allows the use of an empirical distribution. This distribution is derived from

the number of storms that enter into the storm catalog from each calendar year in the rainfall input dataset. Then, during the temporal resampling step,  $k$  is obtained by randomly selecting one of these values. This feature may be useful in regions where storm occurrences exhibit strong clustering (i.e. where there is strong evidence for more storms in some years and fewer in other years for persistent climatological reasons; e.g., Villarini et al., 2013). Section 4.3 examines the impact of this choice on SST results. Other discrete probability distributions, such as the two-parameter negative binomial (Pascal) distribution, can also be used to model clustered count data. RainyDay does not currently use such distributions, since short (typically 10-20 year) remote sensing records may yield poor parameter estimates stemming from the limited number of statistical degrees of freedom.

### 3.5 “Spin-up” of Initial Conditions

A key issue in the modeling of rainfall driven hazards is to adequately represent initial conditions. In many flood and landslide modeling efforts, the most critical of these initial conditions is antecedent soil moisture, while other states such as river baseflow and water table position may also be relevant. Many hydrologic models allow for the specification of such initial conditions, and thus many design storm-based hazard modeling efforts rely on an assumed soil moisture state, such as a typical or fully saturated condition. Such assumed approaches have previously been used with SST (Wright et al., 2014b), and could be combined with the rainfall scenarios generated via RainyDay. This approach has the downside, however, that the true variability antecedent soil moisture is not captured in hazard predictions. This is particularly important in regions in which heavy rainfall does not necessarily occur in the same season as high soil moisture conditions. A second approach that can capture this variability would be to derive a distribution of antecedent soil moisture from previous long-term (ideally multi-decadal) model simulations. Since there can be substantial variation in how soil moisture is represented in different hazard models, ideally the same model would be used for these long-term simulations and for the hazard scenario modeling. RainyDay offers an alternative option, however, in which initial soil moisture can be “spun up” within the hazard model to represent seasonally realistic initial conditions without the need for long-term simulation.

The spin-up procedure is described for a single rainfall scenario. The month of occurrence of the rainfall scenario is identified based on the “parent storm” that created it. Then RainyDay

identifies the set of  $X$ -day periods (where  $X$  is a user-defined spin-up period) preceding all parent storms that occur within a user-defined number of months from the date of occurrence of the parent storm. One of those  $X$ -day periods is randomly selected and pre-pended to rainfall scenario. This scheme helps to ensure that spin-up conditions are reasonable for the given season. It also helps ensure that spin-up conditions have realistic temporal correlations when pre-pended to the rainfall scenario (for example, if there is a historical tendency for several days of moderate rain prior to heavy storms but several days of heavy rain prior to the main storm doesn't have historical precedent, these conditions will be properly represented). It is important to note, however, that the 10 to 20-year records typical of rainfall remote sensing records may not capture the full variability of "true" initial conditions.

This pre-pending procedure creates rainfall scenario output files that are of  $X+t$  day duration. The modeler can then assign an average initial soil moisture condition to initialize each model run, and use the rainfall scenario as input. Soil moisture within the model will then evolve over the spin-up period based on the rainfall (or lack thereof), evapotranspiration, and other model-estimated fluxes. It is important to point out that this spin-up procedure has several limitations. First, it has a storage and computational cost since it can substantially increase the size of the rainfall scenario output files generated by RainyDay and increase the length of each hazard simulation. The importance of these limitations depends on the size of  $A$ , the resolution of the input rainfall dataset, and the computational burden of the hazard model. In Section 4.2, for example, we limit  $X$  to 6 days, for a total rainfall duration of 10 days. This spin-up period is likely sufficient to spin up moisture in the upper soil layers, but not to fully establish baseflow or deeper groundwater flow. The modeler should evaluate the tradeoffs between longer  $X$  and the associated storage and computational costs.

### 3.6 Parametric Rainfall Intensity

Instead of relying on the rainfall intensity derived from a remote sensing input dataset, a user might prefer to use a parametric distribution to impose rainfall depths on the rainfall output scenarios. RainyDay supports this option. The user can supply a  $t$ -hour rainfall depth distribution. This distribution is then applied to the output rainfall scenarios via a normalization procedure that assumes that the supplied distribution corresponds to the annual maximum  $t$ -hour rainfall intensity for a single rainfall grid cell. Rainfall space-time structure is still derived from

the remote sensing data. It should be noted, however, that when the resolution of the input remote sensing dataset is coarse relative to the spatial coverage of the rainfall measurement device upon which the parametric distribution is based (for example, the 16-625 km<sup>2</sup> footprint of many satellite rainfall datasets relative to the 0.1 m<sup>2</sup> sampling area of a single rain gage), this approach may be problematic. This procedure is also problematic in regions where such parametric rainfall distributions might be the synthesis of “mixture distributions” of distinct storm types in which rainfall intensity is intrinsically linked to rainfall space-time structure (e.g. Smith et al., 2011), since RainyDay does not distinguish between different storm types. Currently only the three-parameter generalized extreme value distribution (Walshaw, 2013) is supported, though it would be straightforward to add additional choices.

## **4. Rainfall and Flood Case Studies**

### *4.1 Rainfall IDF*

We generated IDF results for six durations from 3 to 96 hours over a range of  $p_e$  between 0.5 and 10<sup>-3</sup> using RainyDay for single rainfall grid cells in the vicinity of Iowa City, Iowa (Figure 4) using rainfall data from Stage IV (Lin and Mitchell, 2005) and version 7.0 of the Tropical Rainfall Measurement Mission Multi-Satellite Precipitation Analysis (TMPA; Huffman et al., 2010). Stage IV is available through the National Weather Service (NWS) National Center for Environmental Prediction and provides hourly, 4 km resolution rainfall estimates by merging data from the NWS Next-Generation Radar network (NEXRAD; Crum and Alberty, 1993) with rain gages and, in some instances, satellite rainfall estimates. Stage IV has been extensively used in studies of extreme rainfall and flooding. All Stage IV analyses in this paper use data from 2002 to 2014. TMPA merges passive microwave, active radar, and infrared observations from multiple satellites to create a near-global ( $\pm 50^\circ$  latitude) rainfall dataset with 3-hourly, 0.25° (approximately 25 km) resolution. Unless otherwise noted, TMPA analyses this study uses the final “research version” of TMPA from 1998-2014, which includes a monthly rain gage-based bias correction. For the results in Figure 4, and most subsequent analyses in this study,  $A'$  is the rectangular area shown in Figure 3.  $A$  is set to a single rainfall pixel and each run consists of 100 ensemble members ( $N=100$ ), producing 100 estimates for each  $p_e$ . We compare these results with rain gage-based IDFs from NOAA Atlas 14. Atlas 14 uses L-moment regionalization techniques



to combine observations from large number of rain gages. The Atlas 14 analysis for Iowa uses 369 rain gages, many of which have records beginning in the late 19<sup>th</sup> century.

The range of IDF durations shown in Figure 4 emphasize that RainyDay is flexible in terms of the selection of duration  $t$ . RainyDay-based IDF estimates using Stage IV exhibit slight systematic underestimation relative to Atlas 14 across a range of  $p_e$  except for at the 96-hour duration, where there is a close match. RainyDay-based IDF estimates using TMPA, meanwhile, closely match Atlas 14 for high  $p_e$  (except at the 3-hour scale) and underestimates for low  $p_e$  for all durations. Underestimation using RainyDay may be attributed to the mismatch in spatial resolution of the remote sensing data (approximately 16 km<sup>2</sup> for Stage IV and 625 km<sup>2</sup> for TMPA) and the rain gages (approximately 0.1 m<sup>2</sup>). We have refrained from using ARFs to convert the Atlas 14 point IDF estimates into area-averaged IDFs, since the ARF concept has practical and conceptual limitations (see Wright et al., 2014a). Both the slight overestimation of rainfall depth from TMPA (relative to Stage IV) for more frequent events, and the underestimation for more rare events using both datasets, could potentially be explained by conditional bias (i.e. bias that is dependent on rain-rate; Ciach et al. 2000, see Habib et al., 2009 for evidence of conditional biases in TMPA). The convergence between Stage IV-based RainyDay IDFs and Atlas 14 with increasing duration is consistent with both conditional bias and spatial mismatch effects, both of which are known to diminish with increased temporal aggregation. While not definitive, the results in Figure 4 do not clearly point to shortcomings associated with the SST procedure itself.

In order to highlight both the potential for IDF estimation and probabilistic hazard assessment in data-sparse regions using RainyDay with satellite remote sensing, and some of the associated challenges, we compare 24-hour IDF curves generated using RainyDay for various satellite rainfall datasets for the vicinity of Iowa City (Figure 5). This comparison includes two versions of TMPA: the aforementioned final version which includes monthly rain gage-based bias correction, and TMPA-RT, which is produced in near real-time, does not feature bias correction, and runs from 2000-2014. It also includes two versions of the 30-minute resolution, 8 km Climate Prediction Center (CPC) Morphing Technique (CMORPH; Joyce et al., 2004): CMORPH Corrected, which uses a daily rain gage-based bias correction scheme, and CMORPH

Raw, which does not. Finally, it includes the 60-minute, approximately 4 km version of Precipitation Estimation from Remotely Sensed Information Using Artificial Neural Networks Global Cloud Classification System (PERSIANN-GCCS; Sorooshian et al., 2000), which does not use gage-based bias correction. The results in the top panel of Figure 5 show relatively good agreement between point-scale NOAA Atlas 14 IDF's and single-pixel RainyDay-based IDF's for bias-corrected TMPA and PERSIANN-GCCS, particularly considering the spatial sampling mismatch between the remote sensing data and Atlas 14 mentioned previously, while results based on CMORPH Corrected show systematic underestimation.

The middle panel of Figure 5 shows how RainyDay can be used to examine the effect of rain gage-based bias correction on satellite-based IDF estimates. In the case of CMORPH, the Raw version overestimates rainfall intensity at all  $p_e$ , while results for the Corrected version shows that the daily-scale bias correction scheme seems to overcompensate, leading to systematic underestimation. The TMPA-RT also overestimates at all  $p_e$ , though not as severely as CMORPH Raw, while the monthly bias correction scheme used in the final version of TMPA appears to offer superior performance to the daily-scale routine used by CMORPH Corrected. It is not immediately clear why this is the case, particularly since details of the bias correction procedure for CMORPH are not readily available, but relevant considerations include the effect of rainfall detection errors on bias correction (Tian et al., 2007) and the challenge of correcting for conditional biases at short time scales (Wright et al., 2014c). The apparent strong performance of the monthly bias correction is encouraging in the context of Integrated Multi-satellitE Retrievals for GPM (IMERG), a state-of-the-art rainfall dataset that combines various elements from TMPA, CMORPH, and PERSIANN, including TMPA's monthly bias correction (Huffman et al., 2014). The IMERG dataset is not analyzed in this study since the full retrospective dataset is not yet available.

The bottom panel of Figure 5 shows results similar to those in the top panel, but with  $A$  set to a  $0.5^\circ$  by  $0.5^\circ$  (approximately  $2500 \text{ km}^2$ ) box centered on Iowa City. The results demonstrate that RainyDay can easily generate spatially aggregated rainfall IDF curves. This is not achievable using standard gage-based IDF curves without the use of ARFs, which, as previously mentioned,

have been shown to have limitations. We omit an area-averaged gage-based IDF curve from the bottom panel of Figure 5 for this reason.

The results shown in Figure 5 (and also Figure 4) have implications for using RainyDay for IDF and hazard estimation in data-sparse regions using satellite remote sensing. First, there can be substantial differences in extreme rainfall estimates between satellite rainfall datasets, and these differences will propagate through to IDF estimates (and to probabilistic hazard estimates, as will be shown in Section 4.2). Furthermore, while comparison with gage-based IDFs (when available) can be used to understand these differences, spatial sampling mismatches complicate comparisons. Findings may not be transferable across regions since the performance of satellite rainfall retrievals vary with region and latitude (e.g. Ebert et al., 2007) and because the quality of the gage-based bias correction schemes that some of satellite datasets employ will vary regionally with the density of rain gage observations that are available.

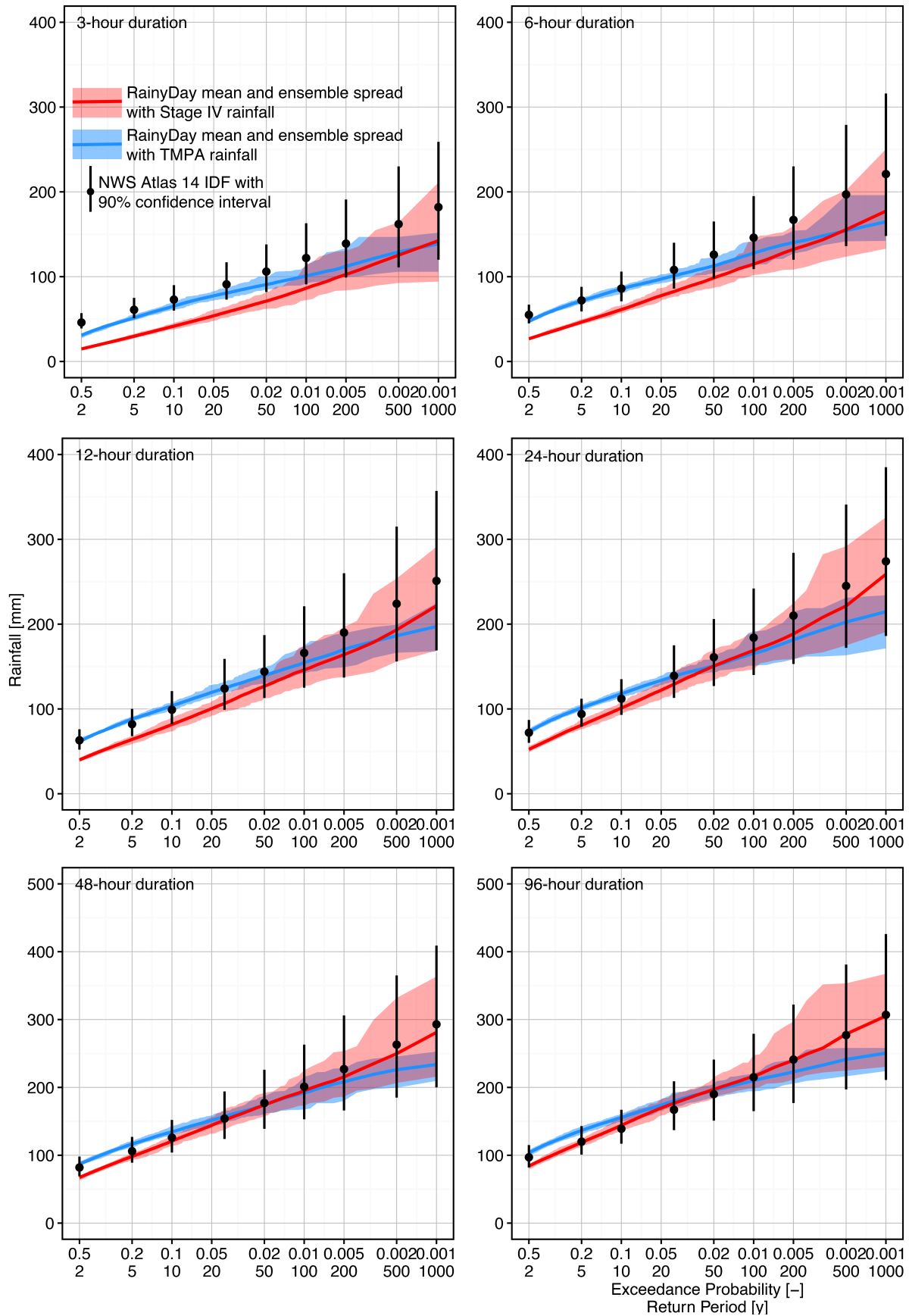


Figure 4: Comparison of IDF curves from Atlas 14 and RainyDay using the Stage IV and TMPA rainfall datasets for 3-, 6-, 12-, 24-, 48-, and 96-hour durations. Shaded areas for RainyDay estimates denote the ensemble spread. Bars on the NOAA Atlas 14 IDF estimates denote the 90% confidence intervals. Key RainyDay parameters:  $m=150$  storms,  $A'=[40^\circ \text{ to } 44^\circ \text{ N}, 90^\circ \text{ to } 96^\circ \text{ W}]$ ,  $A$  is a single rainfall pixel (approximately  $16 \text{ km}^2$  for Stage IV,  $625 \text{ km}^2$  for TMPA),  $N=100$ ,  $T_{max}=1000$ . Spatially-uniform transposition and Poisson-based temporal resampling are selected. Stage IV period of record is 2002-2014, TMPA period of record is 1998-2014. Analyses are restricted to April-November period.

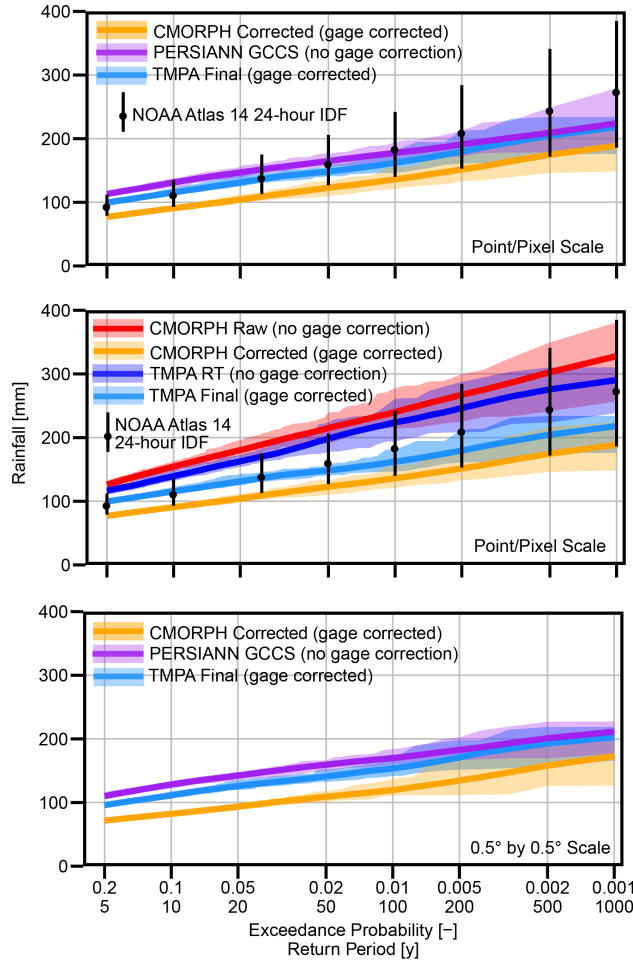


Figure 5: Comparison of IDF curves. Top: 24-hour duration IDF curves at the point scale from NOAA Atlas 14 and at the pixel scale from RainyDay using TMPA Final, PERSIANN-GCCS, and CMORPH Corrected rainfall datasets. Middle: 24-hour duration IDF curves at the point scale from NOAA Atlas 14 and at the pixel scale from RainyDay using TMPA-RT, TMPA Final, CMORPH Raw, and CMORPH Corrected rainfall datasets. Bottom: 24-hour duration IDF curves at the  $0.5^\circ$  by  $0.5^\circ$  scale from RainyDay using TMPA, PERSIANN-GCCS, and CMORPH Corrected rainfall datasets. Shaded areas for RainyDay estimates denote ensemble spread. Bars on the NOAA Atlas 14 IDF estimates denote the 90% confidence intervals. Key RainyDay parameters:  $m=150$  storms,  $A'=[40^\circ \text{ to } 44^\circ \text{ N}, 90^\circ \text{ to } 96^\circ \text{ W}]$ .  $A$  is a single Stage IV rainfall pixel (approximately  $625 \text{ km}^2$  for TMPA,  $64 \text{ km}^2$  for CMORPH,  $16 \text{ km}^2$  for PERSIANN),  $N=100$ ,  $T_{max}=1000$ ,  $t=24$

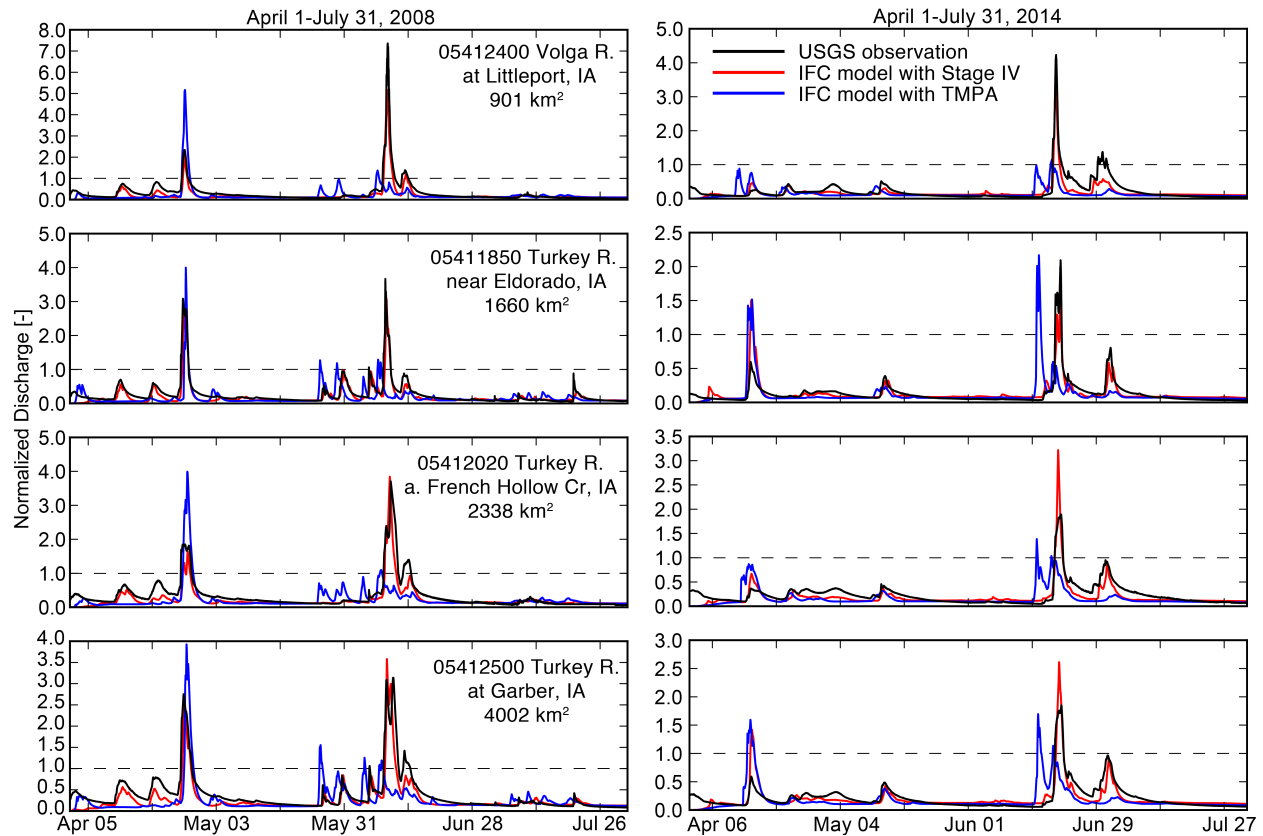
hours . Spatially-uniform transposition and Poisson-based temporal resampling are selected. TMPA Final and CMORPH period of record is 1998-2014, TMPA RT period is 2000-2014, PERSIANN GCCS period of record is 2004-2014. Analyses are restricted to April-November period.

#### 4.2 Flood Frequency Analysis

In this section, we present flood peak frequency analyses for the 4400 km<sup>2</sup> Turkey River watershed in northeastern Iowa using rainfall scenarios from RainyDay as inputs to the Iowa Flood Center (IFC) Model, a calibration-free distributed hydrologic modeling framework designed primarily for multi-scale flood research and application (see Cunha et al., 2012; Demir and Krajewski, 2013; Mantilla and Gupta, 2005; Moser et al., 2015; Small et al., 2013). Moser et al. (2015) provides a detailed model description and Cunha et al. (2012) performed model validations for flood events in Iowa, showing that the performance of the IFC Model is generally comparable to that of the more heavily-calibrated operational NWS SAC-SMA flood forecast model (Burnash, 1995). The model configuration used here is the same that was used by Moser et al (2015). This study aims only to demonstrate basic features of RainyDay for flood hazard analysis and so does not provide detailed discussion of the IFC Model or comparisons with other available platforms. For a discussion of the value of calibration-free, distributed hydrologic models for multi-scale flood modeling, the reader is directed to Wright et al. (2014b) and, in particular, Cunha et al. (2012). The full multi-scale hazard estimation capabilities of SST and RainyDay can, in principle, be harnessed using any distributed hydrologic or mass wasting model, while some of the capabilities can be achieved through the use of lumped models.

A limited set of model hydrograph validation is provided in Figure 6 for the 2008 and 2014 April-July periods, during which major flooding occurred throughout Iowa (see Smith et al., 2013 for a detailed examination of the hydrometeorology of the 2008 flood season). The model is run both with Stage IV and the final (gage-corrected) version of TMPA rainfall. Comparisons with U.S. Geological Survey (USGS) stream gage observations are provided at four locations, with upstream drainage areas ranging from 900 to 4000 km<sup>2</sup>. All hydrographs are normalized by the median annual flood ( $p_e=0.5$ ) to facilitate comparison across watershed scales. Median annual flood estimates are taken from the USGS StreamStats system (<http://water.usgs.gov/osw/streamstats/>). Model performance varies from event to event but there

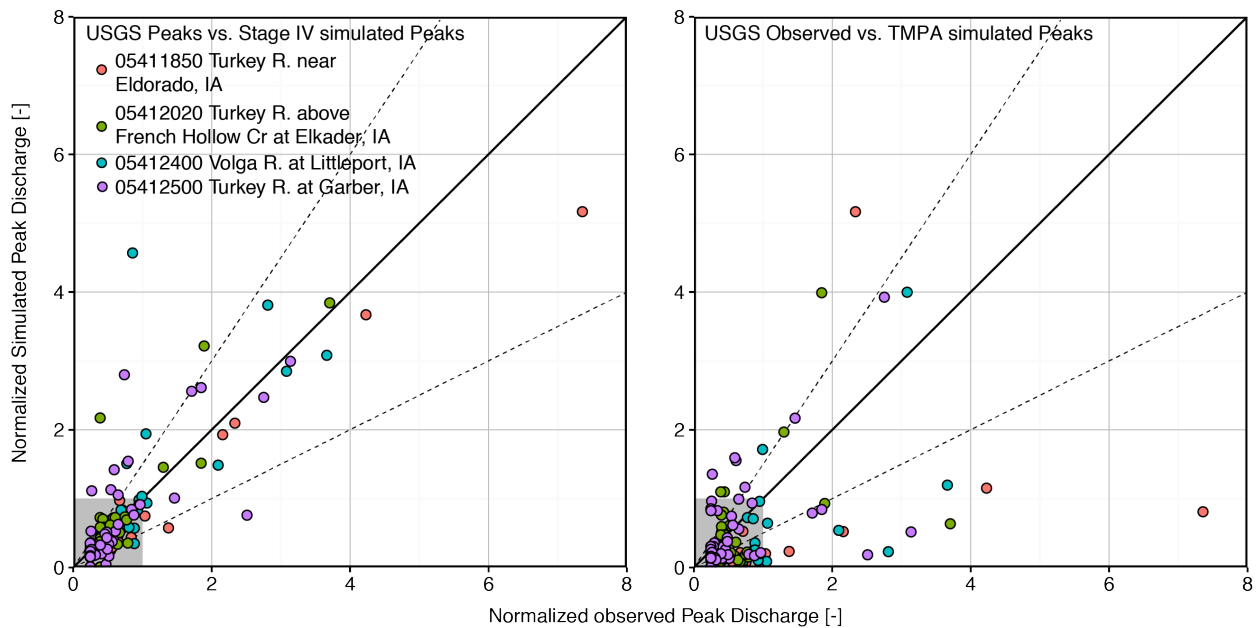
is no clear evidence of systematic bias in the streamflow predictions as a function of event magnitude or drainage area. Predictions based on Stage IV are generally better than TMPA, and in fact several time periods show serious problems with the timing of TMPA-based simulations. In the 2008 flood season, TMPA incorrectly identifies the late April event as the largest for that year, rather than the early June floods.



**Figure 6: IFC model validation for 2008 and 2014 flood seasons (left and right panels, respectively) at four USGS stream gaging sites. Hydrographs are normalized by the median annual flood, which is indicated by dashed horizontal lines.**

We compare observed and simulated flood peaks for the 2008-2014 April-November period (Figure 7). All observed flood peaks that exceed  $100 \text{ m}^3 \text{ s}^{-1}$  are extracted from the four USGS stream gaging records. Then the corresponding flood peaks predicted by the IFC model are extracted from simulated hydrographs based on Stage IV and TMPA rainfall (left panel and right panels of Figure 7, respectively). To allow for modest errors in flood peak timing, a window of 48 hours centered around the observed peak is used to identify the corresponding simulated peaks. All peaks in Figure 7 are normalized by the median annual flood for to facilitate

comparison across basin scales. As a rule of thumb, peaks below the median annual flood can be considered “within bank,” while peaks above the median annual flood can be considered “out-of-bank,” meaning the flood magnitude is large enough to exceed the normal confines of the river channel and spill into the floodplain. The left panel of Figure 7 shows that, while there is modest scatter in the Stage IV-based flood peak simulations, there is no obvious systematic bias with watershed scale or event magnitude. The TMPA-based simulations in the right panel of Figure 7 exhibit greater scatter, generally poor performance, and show some low bias across a range of event magnitudes. While not an exhaustive, the validation shown in Figures 6 and 7 suggests that streamflow prediction accuracy in the IFC model is driven primarily by the accuracy of the input rainfall rather than by model structure, consistent with Cunha et al. (2012), and that the limited accuracy of satellite rainfall inputs, even with gage-based bias correction, can translate into relatively poor model performance.



**Figure 7: Peak discharge validation for 2008-2014 April-November period at four USGS stream gaging stations. All events for which the USGS observations exceeded  $100 \text{ m}^3 \text{ s}^{-1}$  are shown, and peak discharges are normalized by the median annual flood. Simulated peaks using the IFC model with Stage IV (TMPA Final) rainfall inputs are compared with USGS observed peaks in the left (right) panel. Straight black lines indicate 1:1 correspondence, while dashed lines indicate the envelope within which the modeled values are within 50% of observed. Grey boxes in the lower lefthand corners of each panel highlight all events less than the median annual flood.**



We performed IFC model simulations using RainyDay rainfall scenarios developed from both the Stage IV and final gage-corrected TMPA rainfall datasets. For each rainfall dataset, we ran ten ensemble members (i.e.  $N=10$ ), each consisting of 500 rainfall scenarios (i.e.  $T_{max}=500$ ). At any point along the modeled river system, therefore, flood peak exceedance probabilities as low as 0.002 (500-year return period) could be directly derived from the IFC Model predictions. The Stage IV and TMPA-based storm catalogs for the Turkey River include 150 storms, drawn from the April-November rainfall record (2002-2014 for Stage IV, 1998-2014 for TMPA).  $A'$  is an area covering most of Iowa, southwestern Wisconsin, and southeastern Minnesota in the United States.  $t = 96$  hours for all simulations in this section.

We initialize each simulation with a spatially uniform initial soil moisture value found to be typical for the region. Rainfall from a seasonally-based six-day “spin-up” period was then prepended to each 96-hour storm period as per Section 3.5, for a total rainfall input time period of ten days. Spatial variations in both soil moisture and river flow were therefore allowed to develop in each simulation prior to the arrival of the main storm. It should be noted that restricting the rainfall record to April-November, in addition to the lack of snowfall functionality in RainyDay and snowpack functionality in the IFC model, means that snowmelt-driven flooding is not considered in the analyses. In Iowa, snowmelt is generally a minor though non-negligible flood mechanism. We do not evaluate the accuracy of these spin-up soil moisture and river flow, and in fact such evaluation is relatively challenging due to the paucity of long-term soil moisture observation records that would be needed to correlate with river flow. As discussed in Section 3.2, SST and RainyDay facilitates “decoupling” of discharge  $p_e$  from rainfall  $p_e$ . Though not demonstrated explicitly, this decoupling is reflected in the RainyDay-based frequency analyses in this section, in that the role of spun-up initial conditions and rainfall space-time structures can produce discharge  $p_e$  that are different from the  $p_e$  of the input rainfall scenarios.

RainyDay-based frequency analysis results are shown for five subwatersheds of the Turkey River, ranging in drainage area from approximately 460 to 4000 km<sup>2</sup> (Figure 8). Also included in Figure 8 are two types of frequency analyses derived from USGS stream gage observations and taken from Eash et al. (2013) and retrieved from the USGS StreamStats system. The first is developed using standardized methods described in Bulletin 17B (Interagency Advisory

Committee on Water Data, 1982) using the log-Pearson Type III distribution (henceforth referred to as the LP3 distribution) with a regionalized skew coefficient. The second is based on regional regression equations that consider drainage basin area and shape as well as some soil and geological properties. Eash et al. (2013) report 121 years of data for Turkey River at Garber, near Eldorado, and above French Hollow, while 63 years are reported for Volga River at Littleport and 45 years for Turkey River at Spillville. It should be noted that these record lengths refer to “historic record length” described in Section V.B.10 of Bulletin 17B and do not correspond to length of the USGS annual maxima streamflow timeseries available on the USGS National Water Information System (<http://nwis.waterdata.usgs.gov/nwis>), which are much shorter. All available USGS streamflow observations for the five sites are also shown, where  $p_e$  is estimated using the Cunnane plotting position (Cunnane, 1978;  $p_e^i = [i - 0.4] / [X + 0.2]$ , where  $i$  is the rank of the observation and  $X$  is the number of observations). Other common plotting position formulae produce similar results (not shown) and do not alter the conclusions that follow.

For all five locations shown in Figure 8, the SST-based peak discharge estimates using TMPA are higher than those using Stage IV for  $p_e < 0.01$ , generally converging toward the Stage IV estimates as  $p_e$  decreases, and in some cases yielding lower estimates for  $p_e$  less than about 0.005. This is consistent with the rainfall IDF results from RainyDay shown in Figure 4 and are suggestive of conditional biases in the TMPA dataset. This is indeed confirmed in Figure 9, which shows watershed-specific IDF curves for the entire Turkey River watershed from RainyDay using TMPA and Stage IV. The USGS streamflow observations shown in Figure 8 agree reasonably well with the Stage IV-based estimates for  $p_e > 0.5$ , with the exception of the smallest subwatershed, Turkey River at Spillville, where Stage IV produces low peak estimates. For  $p_e < 0.5$ , there is a lack of consistency. For example, Turkey River at Garber shows higher estimates from Stage IV than the streamflow observations, while the reverse is true for Turkey River at French Hollow and near Eldorado. Deviations from the USGS observations do not show a systematic scale dependency.

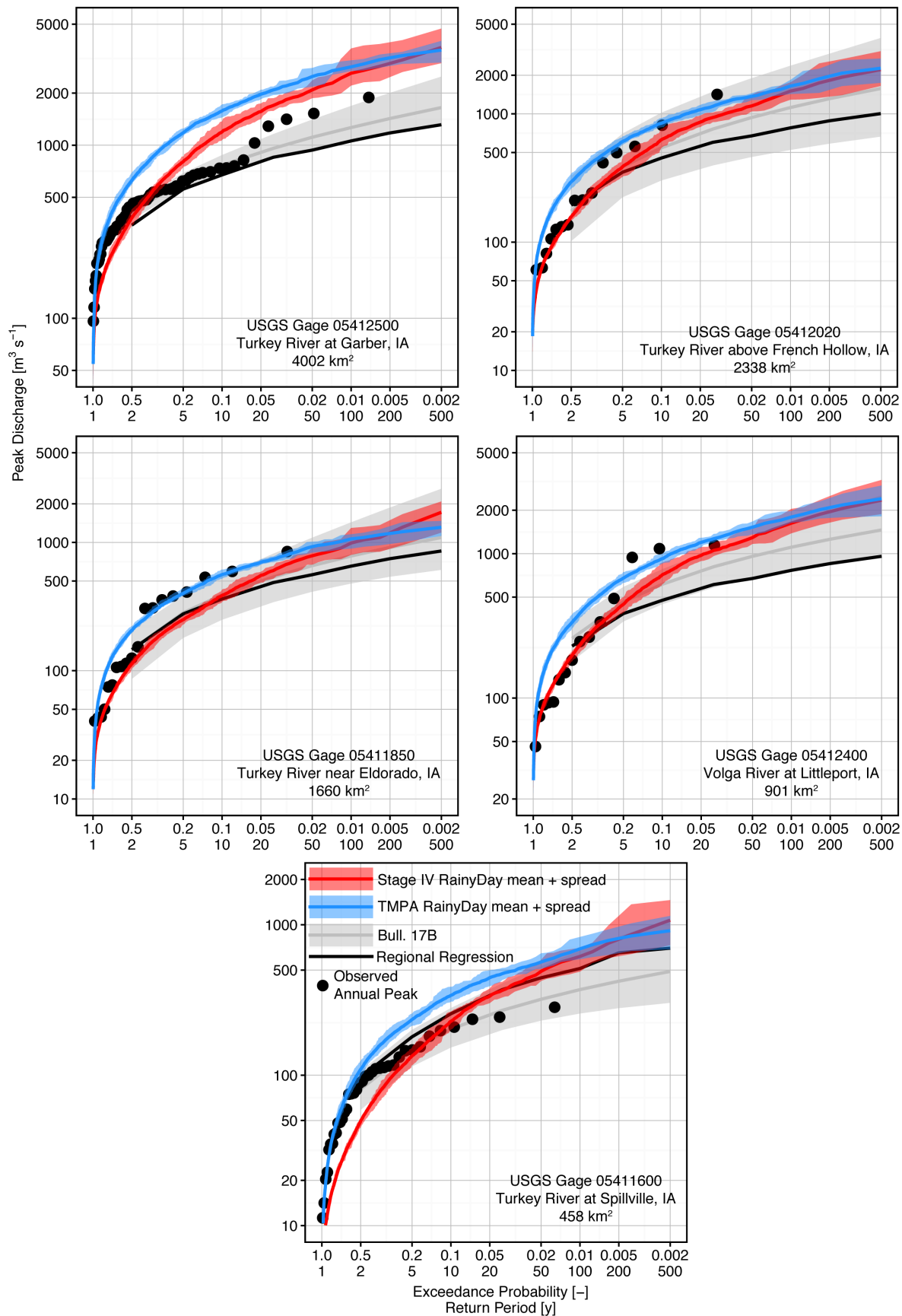
Both RainyDay-based frequency analyses and the USGS streamflow observations are generally higher than the USGS frequency analyses for  $p_e$  less than about 0.2. One exception is the set of USGS observations for Turkey River at Spillville, which are lower than both the RainyDay

estimates and the regional regression but generally consistent with the Bulletin 17B analysis. The regional regression results for Turkey River at Spillville are greater than the USGS regionalized LP3 estimates, while the reverse is true for the four larger subwatersheds. Interestingly, some of the USGS observations fall outside of the 90% confidence intervals of the LP3 analyses for Turkey River near Eldorado, Volga River at Littleport, and Turkey River at Garber. In the case of the latter station, the five most intense floods are near or above the upper 95% confidence bound, a finding that is explored in more detail in the following paragraphs.

It should be noted that with the exception of Turkey River at Garber, the differences between the RainyDay-based analyses are roughly similar in magnitude to the differences between the two different USGS approaches. This fact, along with the underestimation shown by USGS frequency analyses relative to the USGS peak discharge observations at several sites, suggests that the RainyDay-based frequency analyses should not be dismissed out of hand as being too high for low  $p_e$ . In fact, as the next example shows, there is observational evidence that supports the validity of the RainyDay-based results in light of possible nonstationarity in flooding. It should be noted that discharge-based frequency analyses, even in stationary situations with long records, are not necessarily superior to hydrologic modeling methods. Analyses by Smith et al. (2013) suggest that peak discharge measurement errors may be substantial for a recent major flood events in Iowa. The propagation of discharge measurement errors through frequency analysis is poorly understood (e.g., Petersen-Overleir and Reitan, 2009; Petersen-Overleir, 2004; Potter and Walker, 1985). Rogger et al. (2012) reported significant differences between two commonly-used flood frequency analysis approaches for ten small alpine watersheds in Austria, one based on a stream gage-based statistical method and the other on design storm methods combined with a hydrologic model. The latter method produced higher discharge values than the former, and the authors discuss possible explanations and deficiencies in both approaches while concluding that in at least some situations, hydrologic modeling using rainfall inputs will produce superior results.

Of the five USGS stream gage locations shown in Figure 8, only the gage at Garber, Iowa has a long (82-year), unbroken annual peak discharge record. We use this record to better understand the discrepancies between the RainyDay-based results and the USGS frequency analyses from

Eash et al. (2013), and in particular to contrast the methods in the context of potential nonstationarity in flood processes. The top panel of Figure 10 shows that the same RainyDay and USGS frequency analyses shown in Figure 8 for Turkey River at Garber. In this case, however, the USGS observations have been divided into two groups; one for all peaks occurring from 1933 to 1989, and the second for all peaks occurring from 1990 to 2014. The plotting position-based  $p_e$  is recalculated for each group of observations. The 1933-1989 subgroup shows higher discharges than either RainyDay Stage IV or USGS discharges for  $p_e > 0.5$ , and lower discharges for  $p_e$  less than about 0.2. The 1990-2014 subgroup, meanwhile, matches closely with the RainyDay-based frequency analyses with Stage IV.

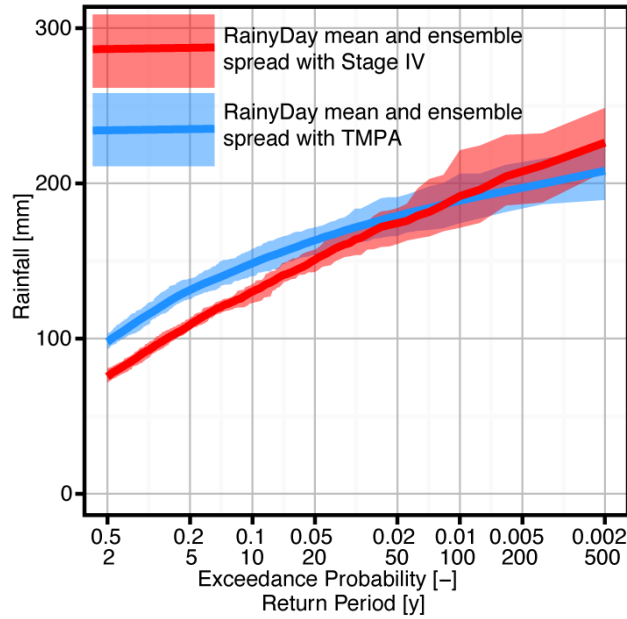


**Figure 8: Peak discharge analyses using RainyDay with Stage IV and TMPA rainfall remote sensing data and the IFC Model, compared against USGS stream gage-based analyses for five subwatersheds of the Turkey River in northeastern Iowa. Shaded areas for RainyDay estimates denote the ensemble spread. Bars on the USGS Bulletin 17B estimates denote the 90% confidence intervals. Confidence intervals are not available for the USGS regional regression. Key RainyDay parameters:  $m=150$  storms,  $A'=[40^\circ \text{ to } 44^\circ \text{ N}, 90^\circ \text{ to } 96^\circ \text{ W}]$ ,  $A$  is the watershed upstream of the USGS streamgage at Garber, IA,  $N=10$ ,  $T_{max}=500$ ,  $t=96$  hours. Spatially-uniform transposition and Poisson-based temporal resampling are selected. Stage IV period of record is 2002-2014, TMPA period of record is 1998-2014. RainyDay Analyses are restricted to April-November period.**

Taken together, this suggests a regime shift toward more extreme flooding since 1990 and a reduction in the magnitude of more average floods. Evidence of this regime shift can be seen in the annual peak time series in the bottom panel of Figure 10. We fit a nonparametric linear regression to the 1933-2014 time series using the nonparametric Theil-Sen estimator (Sen, 1968) and a statistically significant ( $p\text{-value}<0.05$ ) downward trend was found. In contrast, using ordinary least squares, an insignificant upward trend is found over the same period. Thus when the influence of the most extreme values is minimized through nonparametric statistical methods, there is a tendency toward smaller flood peaks over time that is not evident with parametric methods, which are more sensitive to the recent extremes.

The top panel of Figure 10 shows that the period of apparent elevated flood activity is well captured by RainyDay, while the preceding period is not, presumably because the IFC model reflects recent land use changes and because the input rainfall data are relatively recent. In general, whether or not this constitutes a strength or limitation of RainyDay depends on the underlying causation of nonstationary flood activity. If flood nonstationarity results from a climate-driven secular trend in extreme rainfall, then the results from RainyDay using relatively short and recent rainfall remote sensing records should be understood as more “up-to-date” estimates of flood frequency compared to approaches, such as the USGS analyses, that use longer stream gage or rain gage records. The same is true if there is a secular trend in flooding due to urbanization or other land-use changes, so long as these changes are properly incorporated into the hydrologic model. In the case of Iowa, flooding has been shown to be affected by land-use change (Villarini and Strong, 2014) and by climate change (Mallakpour and Villarini, 2015). If, on the other hand, flood or rainfall nonstationarity has a periodic structure due to a slowly-

varying climate mode, then the results from SST may only adequately reflect the true flood frequency for the phase of the mode that overlaps with the remote sensing record. It should also be recognized that a period of higher or lower flood activity at a particular location could result from pure randomness (i.e. in absence of both secular and periodic trends). SST should be relatively robust to this possibility through the sampling storms from a larger region.



**Figure 9: IDF analyses for Turkey River using RainyDay with Stage IV and TMPA rainfall remote sensing data. Shaded areas for RainyDay estimates denote the ensemble spread. Key RainyDay parameters:  $m=150$  storms,  $A' = [40^\circ \text{ to } 44^\circ \text{ N}, 90^\circ \text{ to } 96^\circ \text{ W}]$ ,  $A$  is the  $4400 \text{ km}^2$  watershed upstream of the confluence with the Mississippi River.  $N=100$ ,  $T_{max}=500$ ,  $t=96$  hours, and spatially-uniform transposition and Poisson-based temporal resampling are selected. Stage IV period of record is 2002-2014, TMPA period of record is 1998-2014. Analyses are restricted to April-November period.**

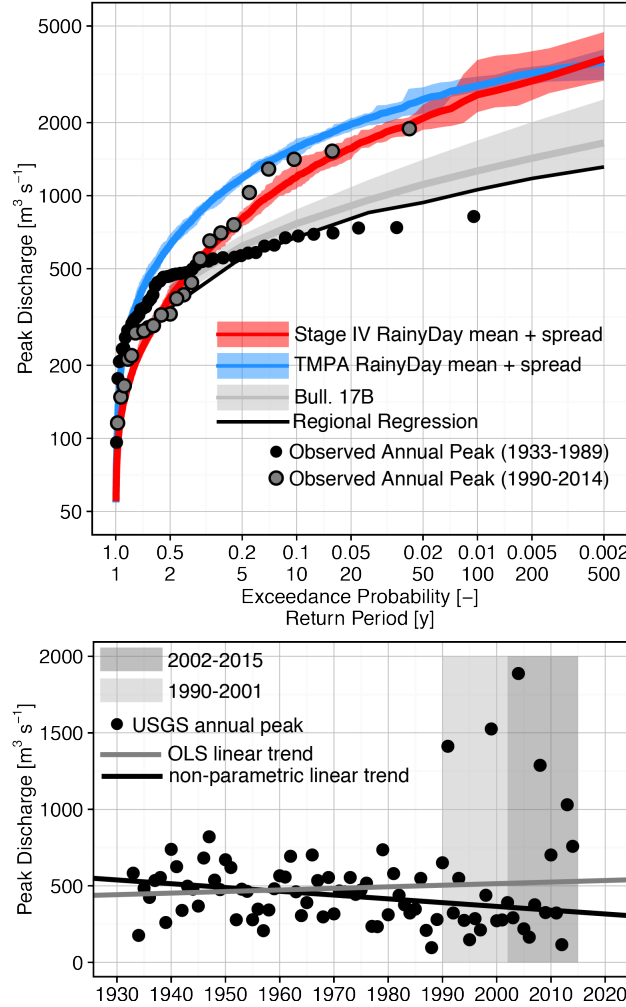


Figure 10: Top panel—four peak discharge analyses for the location of the USGS stream gage at Garber, IA: RainyDay with Stage IV and TMPA rainfall and USGS frequency analyses using regional regression relationships and Bulletin 17B methods. Shaded areas for RainyDay estimates denote the ensemble spread. Bars for the Bulletin 17B-based analysis denote the 90% confidence intervals. Confidence intervals are not available for the USGS regional regression. Bottom panel—annual peak discharge time series for 1932-2014 for the Garber gage. Linear trend lines in the bottom panel use non-parametric Thiel-Sen regression (Sen, 1968) and ordinary least squares (OLS). Key RainyDay parameters:  $m=150$  storms,  $A' = [40^\circ \text{ to } 44^\circ \text{ N}, 90^\circ \text{ to } 96^\circ \text{ W}]$ ,  $A$  is the watershed upstream of the USGS streamgage at Garber, IA,  $N=10$ ,  $T_{max}=500$ ,  $t=96$  hours. Spatially-uniform transposition and Poisson-based temporal resampling are selected. Stage IV period of record is 2002-2014, TMPA period of record is 1998-2014. RainyDay Analyses are restricted to April-November period.



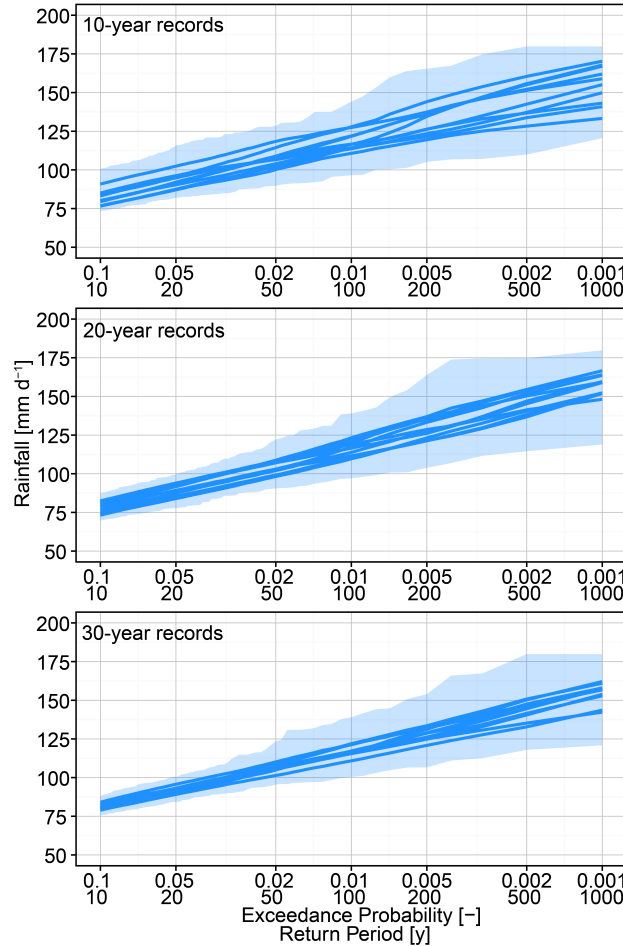
#### 4.3 SST Sensitivity to Record Length and User-defined Parameters

In this section, we examine the sensitivity of SST to the length of the input dataset and to different user-defined parameters and options introduced in Sections 2 and 3. Specific topics that are examined include the optional non-uniform spatial transposition (Section 3.3), empirically-based temporal resampling (Section 3.4) and the size of the transposition domain  $A'$ . In all cases, it should be kept in mind that the specific results pertain to the Iowa study area and may not be generalizable to other locations. The intention is to demonstrate some important concepts and pitfalls associated with RainyDay, and provide a possible framework for assessing performance in different locations and applications.

The core concept behind SST is “space-for-time substitution,” in which storms over a larger region help to inform estimates of rare rainfall in a particular subregion. A common critique of coupling SST with rainfall remote sensing datasets is that such data records are relatively short (approximately 10 to 20 years at time of writing) and thus may not contain sufficient numbers of extreme events at the regional scale to leverage this substitution property and accurately recreate the properties of rare rainfall events. To examine this critique, we turn to a longer dataset: CPC-Unified, a daily rain gage-based gridded rainfall dataset that has a spatial resolution of  $0.25^\circ$  over the conterminous United States (Chen et al., 2008; Xie et al., 2007). Though the spatial and temporal resolution of CPC-Unified is generally insufficient for fine-scale flood modeling, its long record—1948 to present—makes it ideal for evaluating the sensitivity of SST-based IDF estimates to record length. We examined several stationarity measures over the transposition domain  $A'$  (which, as in Section 4.1, roughly encompasses the state of Iowa), including the average number of storm counts per year and the mean, median, and standard deviation of storm rainfall depth (results not shown). None of these measures revealed significant temporal trends, generally consistent with Villarini et al. (2011). This may contradict the apparent flood nonstationarity in the Turkey River watershed discussed in Section 4.2, or may point to land-use change as the predominant source of non-stationarity in Turkey River, but rigorous examination is beyond the scope of this paper.

We use a bootstrapping approach to examine variability in IDF estimates derived from the CPC-Unified data using RainyDay and how this variability evolves as the length of the record

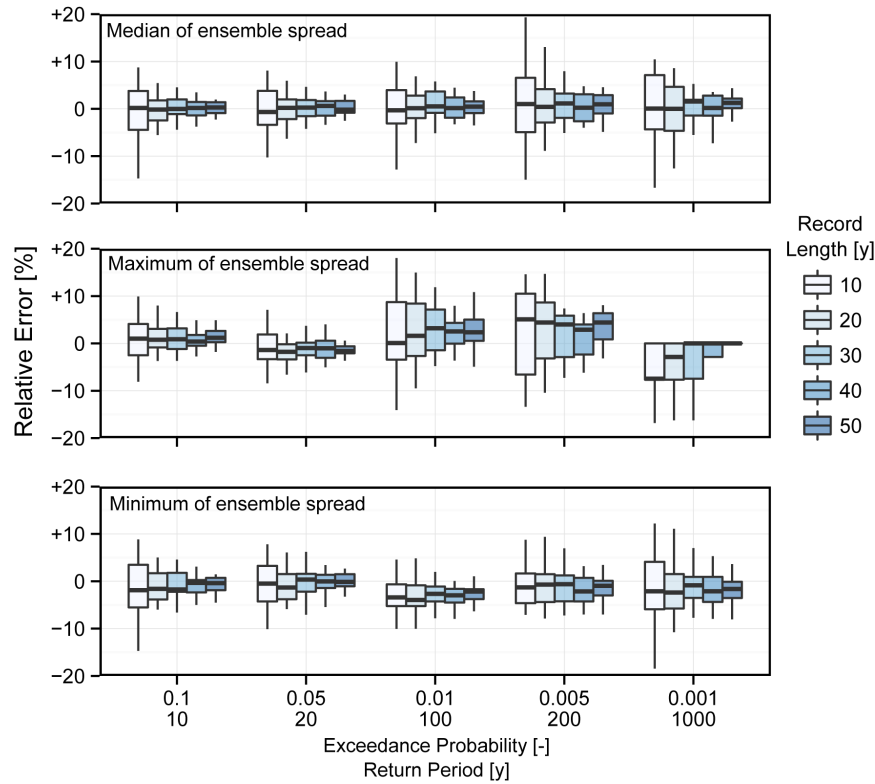
892 increases. All IDF estimates in this section are for 1-day rainfall over averaged over a  $0.5^\circ$  by  
 893  $0.5^\circ$  box. We generate  $n$ -year long input rainfall datasets by randomly selecting  $n$  years of CPC-  
 894 Unified data without replacement from the 1948-2014 period. Each of these datasets is then used  
 895 as the basis for a single run of RainyDay with 100 ensemble members and with  $m = 10n$  (leading  
 896 to  $\lambda=10$  storms per year). We repeat this procedure to create 25 datasets for each value of  $n = 10$ ,  
 897 20, 30, 40, 50 years.



898  
 899 **Figure 11: The effect of the rainfall record length on daily rainfall IDF curves estimated using RainyDay with**  
 900 **the CPC-Unified daily rainfall over Iowa, United States. Each panel shows the ensemble mean (solid lines) for**  
 901 **ten independent runs of RainyDay. The shaded areas denote the maximum spread across the ten runs. Key**  
 902 **RainyDay parameters:  $m=10n$  storms (where  $n$  varies by specified record length),  $A'=[40^\circ$  to  $44^\circ$  N,  $90^\circ$  to  $96^\circ$**   
 903 **W],  $A$  is a  $0.5^\circ$  by  $0.5^\circ$  box,  $N=100$ ,  $T_{max}=1000$ ,  $t=1$  day, spatially-uniform transposition and Poisson-based**  
 904 **temporal resampling. Analyses are restricted to April-November period.**

905  
 906 Substantially more variability is evident in the ensemble mean and spread of the IDF estimates  
 907 using 10 years of CPC-Unified data than using 20 years, while change in variability is negligible

between runs using 20 years and 30 years of data (Figure 11). We also examined the variability of relative deviations in the ensemble IDF means, minima, and maxima from RainyDay between the  $n$ -year runs and IDFs based on the full 67-year dataset (Figure 12). The boxplots show that the majority of the deviations in the  $n$ -year IDF ensemble means, minima, and maxima are less than 10% and that the vast majority are less than 20% for any given  $p_e$ . For most  $p_e$ , there are substantial reductions in deviation when the records increase in length from  $n = 10$  to  $n = 20$  years. The reductions in deviation are less when the record length increases beyond 20 years.



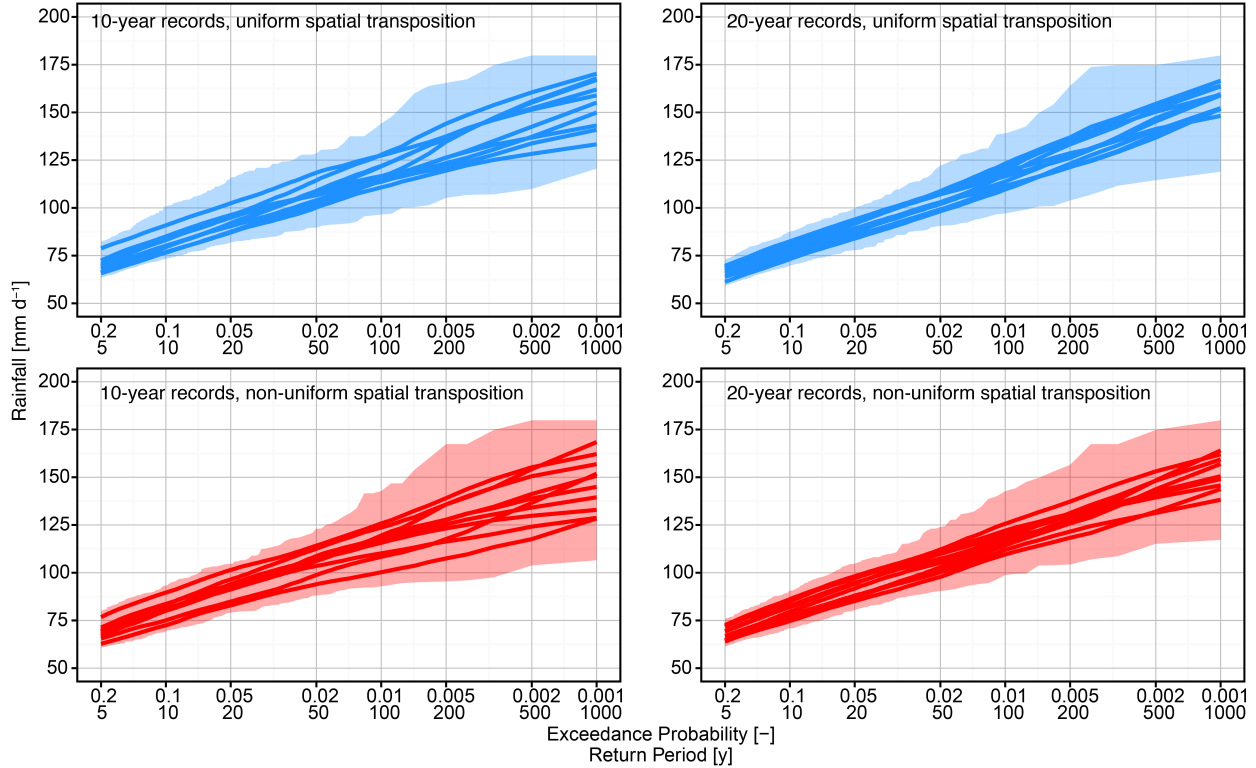
**Figure 12: The effect of rainfall record length on variability in daily rainfall IDF estimated using RainyDay with CPC-Unified data over Iowa, United States for 0.1, 0.05, 0.01, 0.005, and 0.001 exceedance probabilities. Each boxplot shows the variability of a particular rainfall quantity at a given exceedance probability across 25 independent runs of RainyDay. Specific rainfall quantities shown are the ensemble mean (top panel), ensemble maximum (middle panel), and ensemble minimum (bottom panel). Boxes denote the lower and upper quartiles and whiskers indicate the extent of the  $\pm 1.5$  interquartile range. Key RainyDay parameters:  $m=10n$  storms (where  $n$  varies by specified record length),  $A'=[40^\circ \text{ to } 44^\circ \text{ N}, 90^\circ \text{ to } 96^\circ \text{ W}]$ ,  $A$  is a  $0.5^\circ$  by  $0.5^\circ$  box,  $N=100$ ,  $T_{max}=1000$ ,  $t=1$  day, spatially-uniform transposition and Poisson-based temporal resampling. Analyses are restricted to April-November period.**

Unless the intensity of the rainfall inputs is perturbed stochastically, SST-based frequency analyses have an inherent upper bound. This upper bound corresponds to the most intense rainstorm in the storm catalog transposed in such a way that rainfall over  $A$  is maximized. The lack of positive deviations in the ensemble maxima at  $p_e = 10^{-3}$  (middle panel of Figure 12; also in certain realizations shown in Figure 11) show where the SST procedure “encounters” this upper limit.

While the results in this section are by no means exhaustive and the conclusions are specific to the Iowa study region and could vary in different physiographic regions, they nonetheless suggest that concerns over the use of relatively short remote sensing records with SST may be overstated and that remote sensing datasets, many of which are approaching 20 years in length, should provide relatively robust estimates that will improve as these datasets continue to grow in length. This emphasizes the fact that rainfall events that would be considered rare from the perspective of a single location or watershed can occur relatively frequently from a regional perspective. This is qualitatively consistent with the findings of Troutman and Karlinger (2003), who estimate that a flood with  $p_e > 10^{-2}$  occurs on average every 4.5 years at at least one of the 193 USGS stream gage sites in their Puget Sound study region.

A potentially important issue related to short data records in SST, previously mentioned in Section 3.3, can arise if, instead of assuming that the probability of storm occurrence is uniform across the transposition domain, non-uniform spatial transposition is used instead (such as the approach used in Wilson and Foufoula-Georgiou, 1990 or the optional scheme in RainyDay described in Section 3.3). Using the bootstrapping approach with the CPC-Unified dataset described above, visual inspection of storm probability-of-occurrence maps such as the one shown in Figure 3 reveal that there can be substantial variations in the spatial distribution of historical storms when rainfall records are short (results not shown). These variations tend to diminish as the length of record increases, as do their impacts on IDF estimates. More variation is evident in the median IDFs from ten independent runs of RainyDay, for example, using non-uniform transposition than using uniform transposition when  $n=10$  years (Figure 13, left panels). When using non-uniform transposition, variability diminishes when  $n=20$  years and a systematic increase in rainfall intensity for  $p_e > 0.02$ , relative to the uniform transposition case, emerges

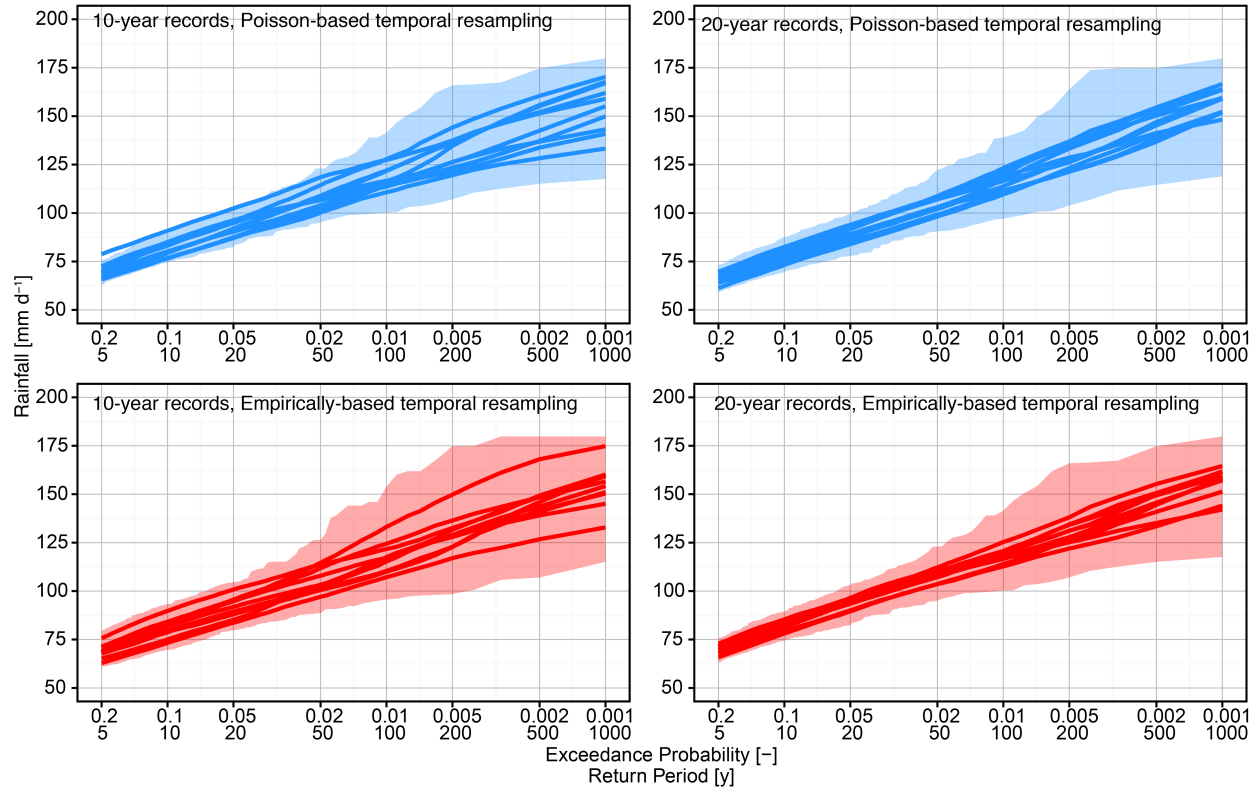
(Figure 13, right panels). Given these results, we recommend that the assumption of uniform transposition be used in the absence of strong physically-based reasoning and observational support for non-uniform transposition. It is possible, however, that this explains the IDF underestimation by RainyDay with Stage IV for high  $p_e$  relative to Atlas 14 shown in Figure 4, where uniform spatial transposition was used.



**Figure 13: The effect of the spatial transposition scheme on daily rainfall IDF curves estimated using RainyDay with the CPC-Unified daily rainfall over Iowa, United States. Each panel shows the ensemble mean (solid lines) for ten independent runs of RainyDay. The shaded areas denote the maximum spread across the ten runs. The specific years that comprise the input dataset vary. Key RainyDay parameters:  $m=10n$  storms (where  $n$  varies by specified record length),  $A'=[40^{\circ}$  to  $44^{\circ}$  N,  $90^{\circ}$  to  $96^{\circ}$  W],  $A$  is a  $0.5^{\circ}$  by  $0.5^{\circ}$  box,  $N=100$ ,  $T_{max}=1000$ ,  $t=1$  day. Poisson-based temporal resampling is used. Analyses are restricted to April-November period.**

As mentioned previously, RainyDay supports either the Poisson-based resampling that has traditionally been used with SST, or an empirical scheme described in Section 3.4. There do not appear to be substantial systematic differences between the results from RainyDay using these two schemes with 10-year records (Figure 14, left panels), but, similar to Figure 13, when 20-year records are used, there is a tendency toward higher rainfall estimates for  $p_e > 0.02$ . Results

may differ in other regions, particularly where temporal clustering of storms is very strong or where rainstorms are very infrequent. It is recommended that the modeler assess clustering using an independent long-term rainfall data source if available, in addition to assessing sensitivity to this option in RainyDay. As with the spatial transposition schemes, the choice of temporal resampling scheme does not appear to have a substantial impact on low  $p_e$  estimates.



**Figure 14: The effect of the temporal resampling scheme on daily rainfall IDF curves estimated using RainyDay with the CPC-Unified daily rainfall over Iowa, United States. Each panel shows the ensemble mean (solid lines) for ten independent runs of RainyDay. The shaded areas denote the maximum spread across the ten runs. The specific years that comprise the input dataset vary. Key RainyDay parameters:  $m=10n$  storms (where  $n$  varies by specified record length),  $A'=[40^{\circ}$  to  $44^{\circ}$  N,  $90^{\circ}$  to  $96^{\circ}$  W],  $A$  is a  $0.5^{\circ}$  by  $0.5^{\circ}$  box,  $N=100$ ,  $T_{max}=1000$ ,  $t=1$  day. Spatially uniform transposition is used. Analyses are restricted to April-November period.**

We also examine the sensitivity of RainyDay results to the size of  $A'$  (Figure 15). To do so, we run RainyDay for various square domains ranging from  $1^{\circ}$  by  $1^{\circ}$  up to  $10^{\circ}$  by  $10^{\circ}$ , while holding  $A$  fixed at a  $0.5^{\circ}$  by  $0.5^{\circ}$  box. Then the evolution of rainfall intensity is examined for a range of  $p_e$  as a function of  $A'$ . This is repeated for a several different record lengths and for two values of

995  $\lambda$ . Interestingly, while there is a general tendency for intensity estimates to stabilize as  $A'$  grows,  
996 the behavior is not asymptotic (though roughly so for  $n=68$  years). The high exceedance  
997 probability estimates ( $p_e=0.5$ ) tend to be stable over a large range of  $A'$  and then decrease for  
998 very large values, due to the tendency for synthetic years to be created in which no storm is  
999 transposed directly over  $A$ . This is the root of potential low biases mentioned in Step 2 of the  
1000 SST procedure described in Section 2. However, Figure 15 demonstrates that this tendency for a  
1001 decrease in intensity estimates for large  $A'$  extends to smaller  $p_e$  values as well, and that there is a  
1002 critical value of  $A'$  at which the estimated intensity is roughly maximized. This critical value  
1003 appears to vary more by the particular period of record than by the length of record. For  
1004 example, the 20-year record from 1976-1995 yielded a critical value of  $A'$  that is lower than the  
1005 critical value from 20-year record from 1996-2015. This points to the fact that the existence and  
1006 number of major storms within  $A'$  during the record period is very important (Wright et al.,  
1007 2014b reached the same conclusion through different means).

1008  
1009 These results also indicate that increasing  $m$  (thus increasing  $\lambda$ ) can mitigate the reduction in  
1010 estimated intensity for values of  $A'$  larger than the critical value. This result suggests that, if the  
1011 modeler is interested in hazard estimation across a range of  $p_e$ , he or she should choose a  
1012 relatively large  $m$ . A diagnostic framework within the RainyDay software to identify this critical  
1013 value of  $A'$  for a given value of  $m$  (or vice versa) for different  $p_e$  would be useful but does not  
1014 currently exist.

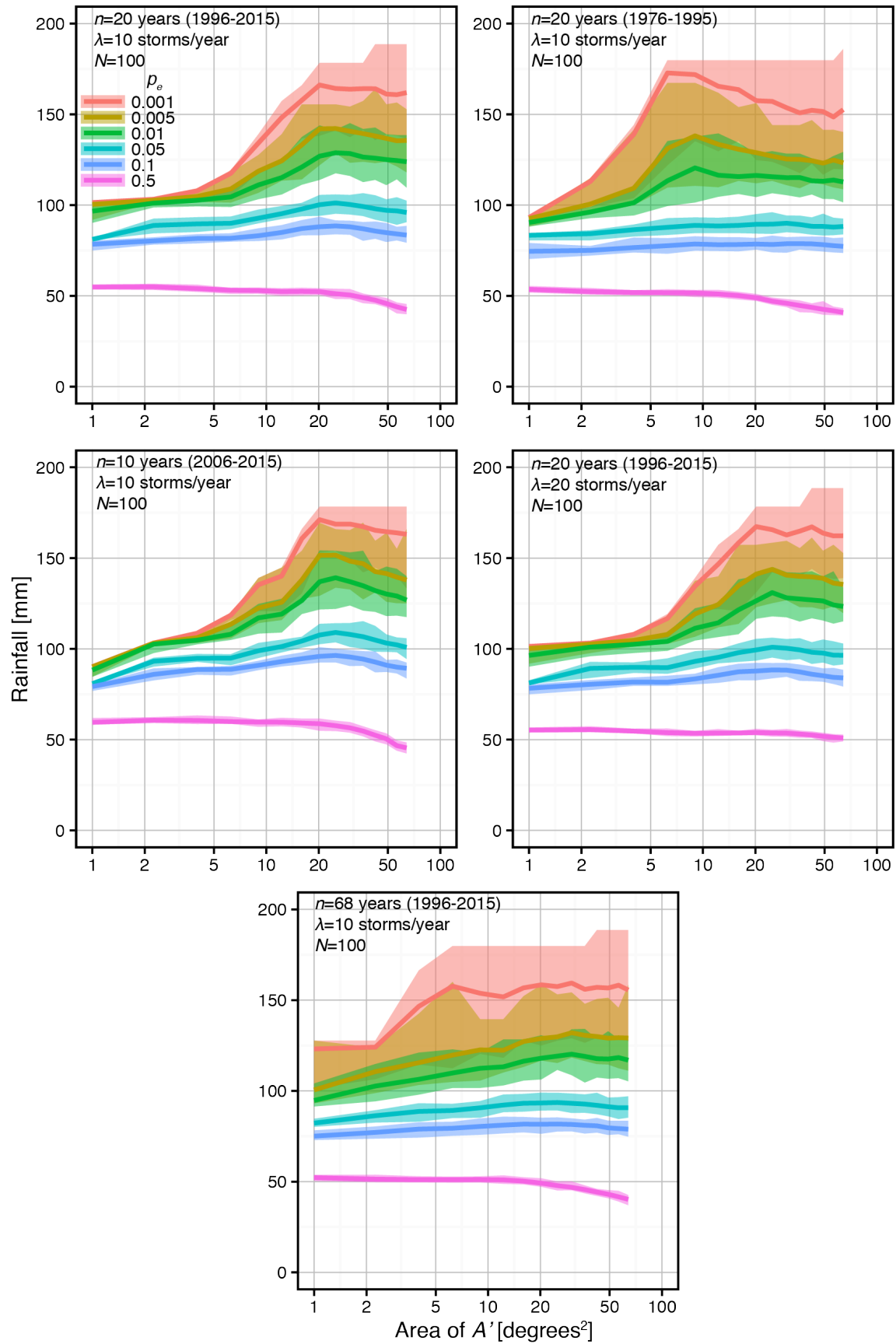




Figure 15: The effect of the size of the transposition domain  $A'$  on daily rainfall IDF curves estimated using RainyDay with the CPC-Unified daily rainfall over Iowa, United States using a range of record lengths. Key RainyDay parameters:  $m=10n$  storms (where  $n$  varies by specified record length),  $A'$  is a square of varying size,  $A$  is a  $0.5^\circ$  by  $0.5^\circ$  box,  $N=100$ ,  $T_{max}=1000$ ,  $t=1$  day, spatially-uniform transposition and Poisson-based temporal resampling. Analyses are restricted to April-November period.

## 5. Discussion and Conclusions

In this paper we introduce RainyDay, a Python-based platform that couples rainfall remote sensing data with a technique known as Stochastic Storm Transposition (SST) that effectively “lengthens” the extreme rainfall record through temporal resampling and spatial transposition of observed rainstorms. It produces probabilistic extreme rainfall scenarios that include realistic estimates of rainfall duration, intensity, and space-time structure that can be used for probabilistic flood and landslide hazard and risk assessment at a wide range of scales.

The SST technique, as implemented in RainyDay, has two important features that distinguish it from IDF and design storm methods for describing the relationships between the intensity, duration, and structure of extreme rainfall. First, it leverages the detailed picture of rainfall space-time structure offered by ground-based radar or satellite-based sensors. This structure can play an important role in landslides and floods because the variability in the concentration and intermittency of extreme rainfall in space and time can lead to a complex and diverse spectrum of hazard response. This structure is difficult to measure using rain gages due to the high gage densities and sampling rates required, and so rain gage-based methods for analysis of rainfall-driven hazards, such as IDF relations and design storm methods, typically neglect this higher-order variability. The reader is directed to Wright et al. (2014b) for a deeper examination of this feature in the context of urban flood hazards.

The second important feature of RainyDay is that, because of the near-global coverage of satellite rainfall datasets, it is possible to generate realistic representations of extreme rainfall in remote or poorly-instrumented regions where rain gage or stream gage records are lacking. Such regions are common even in wealthy nations and are ubiquitous in developing countries, many of which are characterized by rapidly-growing exposure to rainfall-driven hazards due to urbanization and climate change. The authors are not aware of other approaches that offer the

ability to generate realistic rainfall inputs for probabilistic hazard modeling nearly anywhere on the globe with minimal computational effort.

Despite the advantages that SST and RainyDay offer over some other methods for assessing rainfall-driven hazards (e.g. design storms, discharge frequency analysis), a number of limitations and unanswered questions remain. Perhaps the biggest limitation to coupling SST with rainfall remote sensing, and to remote sensing applications more broadly, is the uncertain accuracy of the input rainfall data. Significant efforts have been made to better understand and minimize the errors in remote sensing estimates of rainfall, both from satellites (e.g. Petty and Krajewski, 1996; Tian and Peters-Lidard, 2007; Tian et al., 2009) and from ground-based radar (e.g. Villarini and Krajewski, 2010). Such studies demonstrate that remote sensing estimates can vary significantly from reference observations in terms of rainfall intensity bias and differentiation between rainy and non-rainy areas, with important implications for hazard applications. In the case of satellite-based rainfall estimates, heterogeneities in the underlying land or water surfaces can be difficult to distinguish from variations in cloud and rainfall properties (e.g. Ferraro et al., 2013), while both ground-based radar and space-based sensors tend to suffer in mountainous areas due to dramatic variations in rainfall physical properties over short time and length scales. Furthermore, the spatial and temporal resolution of remote sensing estimates, particularly from satellites, can be too coarse for modeling at very small scales, especially in urban areas and fast-responding mountain or desert catchments where surface runoff generation from intense, short-duration rainfall on sub-hourly, sub-kilometer scales can be a key driver of hazards. The uncertainties associated with rainfall remote sensing data pose serious challenges for flood or landslide forecasting and monitoring, which require accurate rainfall estimates in real-time. These issues may be somewhat less critical in the SST framework or in long-term hazard assessment more generally, since the rainfall estimates need only have fidelity in the statistical sense. SST will be somewhat robust to random errors in rainfall data, as the underestimation of rainfall intensity from some storms in the storm catalog can be compensated by overestimation of rainfall intensity from others. In contrast, SST is not robust to systematic rainfall biases, as demonstrated in several examples in this paper. IMERG, NASA's newest satellite multi-sensor dataset, will feature improved accuracy and relatively high

resolution ( $0.1^\circ$ , 30-minute), addressing some of these issues once the full retrospective dataset from 1998-present becomes available.

In the case of flood hazard modeling using SST, a practical upper limit on the size of the area of interest  $A$  can arise. As mentioned in Sections 2.1 and 3.3, the sizes of  $A$  and  $A'$  can be limited due to the challenges posed by transposition in the presence of complex terrain features. Furthermore, as  $A$  becomes larger, the rainfall duration  $t$  needed to properly model hazard response becomes longer. While RainyDay does not restrict the choice of  $t$ , practical limitations exist. In large watersheds, floods are usually the result of specific space-time arrangements of multiple distinct storm systems over the span of perhaps a week and up to months, often linked to persistent large-scale atmospheric disturbances. One could specify a long  $t$  (a month, for example) in RainyDay to “capture” all of these storm systems within a single storm catalog entry. Such long  $t$ , however, means there could only be relatively few entries in the storm catalog, given the limited record length of the input dataset. Such an approach would be constrained by the few space-time configurations of these storm systems that were actually observed, while other non-observed configurations are hypothetically possible. A tradeoff thus emerges as  $A$  (and thus  $t$ ) increase relative to the area of the transposition domain  $A'$ . If  $A$  is a large fraction of  $A'$ , then there is little opportunity to leverage the “space-for-time” substitution that is at the core of the SST approach. If the user instead decides to increase the size of  $A'$ , he or she must ensure that this transposition is performed in a realistic manner. This effectively precludes modeling of regions that approach continental scales. The maximum scale at which SST can be feasibly used is an open question with no simple answer. It should be noted that IDF and design storm methods face similar and perhaps even more acute limitations in terms of an upper area limit, though for different reasons (e.g. conceptual and practical shortcomings of point-based temporal rainfall distributions and area reduction factors).

As mentioned in Section 4.3, a common critique of the methodology presented in this study is that the relatively short remote sensing records may not contain enough truly extreme rainfall events. Sensitivity to record length is not unique to SST; frequency estimates of rare hazards will be driven by the largest several events in the historical record, regardless of the chosen analysis technique. The results in Section 4.3 demonstrate that this concern may be somewhat

exaggerated in the case of SST since very extreme rainfall events that are considered rare from a local viewpoint can occur much more frequently when viewed regionally. Like more commonly-used regionalization techniques, SST helps to leverage this fact to improve hazard analysis. As the rainfall remote sensing record grows, the robustness of estimates produced by SST and RainyDay should increase as additional extreme storms are observed (and as their accuracy improves due to technological advances). Estimates of rainfall intensity will improve more per unit of time using SST than using point-based techniques due to SST's regional nature, while new patterns of rainfall space-time structure will add to the realism of SST-based flood and landslide hazard estimates since a broader spectrum of hazard outcomes will be possible. RainyDay makes such updating simple, while IDF databases and design storm methods are generally updated through slow and costly procedures (Y. Zhang, personal communication, May 14, 2015).

As highlighted in Section 4.2, SST and RainyDay have important features in the context of nonstationary hazards. Extreme rainfall scenarios from RainyDay are generally based on more recent observations than existing rain gage or stream gage-based frequency analyses such as Atlas 14 IDF relations, which contain older records that may not be representative of the current state of the climate. In this respect, hazard analysis based on RainyDay can be understood as a relatively current "snapshot" based on recent climate. The performance of RainyDay is very dependent on major storms having occurred one or more times within the transposition domain, however, meaning that spatial transposition is not a perfect remedy for short data records. Furthermore, if the rainfall remote sensing record deviates significantly from the true long-term properties of extreme rainfall over the region of interest due to random chance, decadal-scale climate variability, or systematic measurement bias, then caution must be taken when using RainyDay. It can be challenging in practice to diagnose such nonstationarities and biases due to a lack of long-term independent observational data, particularly in remote or underdeveloped regions. Meanwhile, as discussed in Wright et al. (2014b), combining SST (or other rainfall-based approaches, e.g. Cunha et al., 2011) with a distributed hazard model allows the analyst to incorporate changes in land use and land cover into hazard estimates.

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